

A QUINTUPLE LAW FOR MARKOV-ADDITIVE PROCESSES WITH PHASE-TYPE JUMPS

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Abstract

We consider a Markov-additive process (MAP) with phase-type jumps, starting at zero. Given a positive level u , we determine the joint distribution of the undershoot and overshoot of the first jump over the level u , the maximal level before this jump, the time of attaining this maximum, and the time between the maximum and the jump. The analysis is based on first passage times and time reversion of MAPs. A marginal of the derived distribution is the Gerber-Shiu function, which is of interest to insurance risk. Several examples serve to compare the present result with the literature.

Keywords: Markov additive process; undershoot; surplus; ruin; Gerber-Shiu function

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1. Introduction

In recent literature on insurance risk, the undershoot of a variety of stochastic processes has been of considerable interest under the name of "surplus prior to ruin". Mostly it has been investigated within the framework of the so-called Gerber-Shiu (discounted penalty) function, which describes the joint distribution of the time to ruin, the surplus prior to ruin, and the deficit at ruin. Following are some examples of stochastic processes for which results are available: the compound Poisson model [16], its perturbed [15, 12, 22] and Markov-modulated [4, 25] versions, the Lévy risk process [14], the fluid flow model [9, 1], the Sparre Andersen model with Erlang inter-claim times [17] as well as its perturbed version [20]. This list is by no means exhaustive, but one can already see that all these models are special instances of Markov-additive processes. The almost universal approach in the literature is to derive some (defective) renewal equations, starting from a set of differential equations that can be obtained via Itô's formula or the infinitesimal generator of the surplus process (see discussion to [21]).

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From another perspective, a recent paper by Doney and Kyprianou [13] provides a formula for the joint distributions of the space–time positions of overshoots and undershoots for Lévy processes (theorem 3 therein, with example 8 dedicated to insurance risk). Their approach of analysis is more along the classical lines of fluctuation theory, using the ladder height process and time–reversal. For the case of Lévy processes with phase–type jumps, the ladder process is derived explicitly in [27]. Related results on MAPs can be found in [19].

The present paper aims to apply the classical approach of ladder heights and time–reversal to the class of Markov–additive processes (MAPs) with phase–type jumps. Its main result is an explicit formula for the measure

$$\mathbb{E}(e^{-\gamma\tilde{G}_{\tilde{\tau}(u)}-\gamma^*(\tilde{\tau}(u)-\tilde{G}_{\tilde{\tau}(u)})}, \tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy)$$

where $\tilde{\mathcal{X}} = (\tilde{X}_t : t \geq 0)$ is the level process of a MAP, $\tilde{\tau}(u)$ is the first passage time over some level $u > 0$, the $\gamma, \gamma^* \geq 0$ are time discounting factors (which can be seen as the variables for the Laplace transforms of $\tilde{G}_{\tilde{\tau}(u)}$ and $\tilde{\tau}(u) - \tilde{G}_{\tilde{\tau}(u)}$), $\tilde{M}_{\tilde{\tau}(u)}$ is the supremum of $\tilde{\mathcal{X}}$ before the passage time $\tilde{\tau}(u)$, $\tilde{G}_{\tilde{\tau}(u)}$ is the time of attaining this supremum, and $U_{\tilde{\tau}(u)} := u - \tilde{X}_{\tilde{\tau}(u)-}$ resp. $O_{\tilde{\tau}(u)} := \tilde{X}_{\tilde{\tau}(u)} - u$ denote the undershoot resp. the overshoot at time $\tilde{\tau}(u)$. This result will provide all the information (and more) that is usually contained in the Gerber–Shiu function.

Of fundamental use in this paper will be the recent determination of the Laplace transform of first passage times for MAPs as given in [11]. A second pillar of the present work is theorem 2.5 in [4] (see also [2], theorem 3.1, for a queueing context), which yields a relation between an occupation measure and the ladder height via time–reversal. The original result was presented in the framework of the Markov–modulated compound Poisson model.

In the following section we shall collect all the necessary preliminary results that we will need later on. In particular we simplify the results from [11] for the special kind of MAPs that we employ in this paper. Section 3 contains the main result with some corollaries. The final section 4 presents applications to insurance risk, in particular the Gerber–Shiu function, and compares the present result with existing ones in the literature.

2. Preliminaries

2.1. Markov–additive processes with phase–type jumps

Let $\tilde{\mathcal{J}} = (\tilde{J}_t : t \geq 0)$ be an irreducible Markov (jump) process with finite state space \tilde{E} and infinitesimal generator matrix $\tilde{Q} = (\tilde{q}_{ij})_{i,j \in \tilde{E}}$. We call \tilde{J}_t the phase at time $t \geq 0$ (another common name is regime). Define the real–valued process $\tilde{\mathcal{X}} = (\tilde{X}_t : t \geq 0)$ as evolving like a Lévy process $\tilde{\mathcal{X}}^{(i)}$ with parameters $\tilde{\mu}_i$ (drift), $\tilde{\sigma}_i^2$ (variation), and $\tilde{\nu}_i$ (Lévy measure) during intervals when the phase equals $i \in \tilde{E}$. Whenever $\tilde{\mathcal{J}}$ jumps from a state $i \in \tilde{E}$ to another state $j \in \tilde{E}$, $j \neq i$, this may be accompanied by a jump of $\tilde{\mathcal{X}}$ with some distribution function \tilde{F}_{ij} . Then the two–dimensional process $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$ is called a Markov–additive process (or shortly MAP). A MAP can also be defined by the following property (see [5], section XI.2a): $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$ is a Markov process such that

$$\mathbb{E}(f(\tilde{X}_{t+s} - \tilde{X}_t)g(\tilde{J}_{t+s})|\mathcal{F}_t, \tilde{J}_t = i) = \mathbb{E}(f(\tilde{X}_s)g(\tilde{J}_s)|\tilde{X}_0 = 0, \tilde{J}_0 = i) \quad (1)$$

holds for all $s, t > 0$ and $i \in \tilde{E}$, where f and g are measurable functions and $(\mathcal{F}_t : t \geq 0)$ denotes the canonical filtration of $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$. For a textbook introduction to MAPs see [5], chapter XI.

Denote the indicator function of a set A by 1_A . We assume that the Lévy measures $\tilde{\nu}_i$ have the form

$$\tilde{\nu}_i(dx) = \lambda_i^+ 1_{x>0} \alpha^{(ii)+} \exp(T^{(ii)+}x) \eta^{(ii)+} dx + \lambda_i^- 1_{x<0} \alpha^{(ii)-} \exp(-T^{(ii)-}x) \eta^{(ii)-} dx$$

for all $i \in \tilde{E}$, where $\lambda_i^\pm \geq 0$ and $(\alpha^{(ii)\pm}, T^{(ii)\pm})$ are representations of phase–type distributions without an atom at 0. The $\eta^{(ii)\pm} := -T^{(ii)\pm} \mathbf{1}$ are called the exit vectors, where $\mathbf{1}$ denotes a column vector of appropriate dimension with all entries being 1. This means that the jump process induced by the Lévy measure $\tilde{\nu}_i$ is compound Poisson with jump sizes of a doubly phase–type distribution. Denote the order of $PH(\alpha^{(ii)\pm}, T^{(ii)\pm})$ by m_{ii}^\pm . Further write $\lambda_i := \lambda_i^+ + \lambda_i^-$.

Likewise, let p_{ij}^+ (resp. p_{ij}^-) denote the probability that a positive (resp. negative) jump is induced by a phase change from $i \in \tilde{E}$ to $j \in \tilde{E}$, and assume that these jumps have a $PH(\alpha^{(ij)\pm}, T^{(ij)\pm})$ distribution without an atom at 0. Note that $p_{ij}^+ + p_{ij}^- \leq 1$ for all $i, j \in \tilde{E}$. Let m_{ij}^\pm denote the order of $PH(\alpha^{(ij)\pm}, T^{(ij)\pm})$ and define $\eta^{(ij)\pm} := -T^{(ij)\pm} \mathbf{1}$.

We shall exclude the case of $\tilde{\mu}_i = \tilde{\sigma}_i^2 = 0$ for any phase $i \in \tilde{E}$, which would govern the

zero process or a pure Lévy measure. This avoids the awkward case of a non-unique time $G_{\tilde{\tau}(u)}$ of attaining the supremum of $\tilde{\mathcal{X}}$ before the passage time $\tilde{\tau}(u)$.

The class of Markov-additive processes with these assumptions of phase-type jumps is dense within the class of all MAPs, see [6], proposition 1. The main advantage of the restriction on the jump distribution is the possibility of transforming the jumps into a succession of linear pieces of exponential duration (each with slope 1 or -1) and retrieving the original process via a simple time change, see [3], section 3.4.4, or [8]. The path transformation is illustrated in figures 1 and 2, where figure 1 shows a typical path of the original level process $\tilde{\mathcal{X}}$ and figure 2 shows the transformed path \mathcal{X} .

FIGURE 1: Typical path of $\tilde{\mathcal{X}}$

In more exact terms, the transformation is done in the following way. Without the jumps, the Lévy process $\tilde{\mathcal{X}}^{(i)}$ during a phase $i \in \tilde{E}$ is either a linear drift (of positive or negative slope $\tilde{\mu}_i \in \mathbb{R}$) or a Brownian motion (with parameters $\tilde{\sigma}_i > 0$ and $\tilde{\mu}_i \in \mathbb{R}$). Considering this MAP (without the jumps) we can partition its phase space \tilde{E} into the subspaces E_p (for positive drifts), E_σ (for Brownian motions), and E_n (for negative drifts). Then we introduce two new phase spaces

$$E_\pm := \{(i, j, k, \pm) : i, j \in E_p \cup E_\sigma \cup E_n, 1 \leq k \leq m_{ij}^\pm\} \quad (2)$$

to model the jumps. Define now the enlarged phase space $E := E_+ \cup E_p \cup E_\sigma \cup E_n \cup E_-$

FIGURE 2: Typical path of \mathcal{X}

and let $E_c := E_p \cup E_\sigma \cup E_n$ denote the subspace of E that contains all phases under which the real time movements are continuous.

We define the modified MAP $(\mathcal{X}, \mathcal{J})$ over the phase space E as follows. Set the phase-dependent parameters as $(\mu_i, \sigma_i^2, \nu_i) := (\tilde{\mu}_i, \tilde{\sigma}_i, \mathbf{0})$ for $i \in E_c$ and $(\mu_h, \sigma_h^2, \nu_h) := (\pm 1, 0, \mathbf{0})$ for $h \in E_\pm$, i.e. $h = (i, j, k, \pm)$ where $i, j \in E_c$ and $1 \leq k \leq m_{ij}^\pm$. This leads to the cumulant functions $\psi_h(\alpha) = \pm\alpha$ for $h \in E_\pm$ and

$$\psi_i(\alpha) = \begin{cases} \mu_i \alpha, & i \in E_p \cup E_n \\ \frac{1}{2} \sigma_i^2 \alpha^2 + \mu_i \alpha, & i \in E_\sigma \end{cases} \quad (3)$$

where the earlier exclusion of a phase $i \in \tilde{E}$ with $\tilde{\mu}_i = \tilde{\sigma}_i = 0$ yields $\mu_i > 0$ for $i \in E_p$ and $\mu_i < 0$ for $i \in E_n$. The modified phase process \mathcal{J} is determined by its generator matrix $Q = (q_{ij})_{i,j \in E}$. For this the construction above yields

$$q_{ih} = \begin{cases} \tilde{q}_{ii} - \lambda_i, & h = i \in E_c \\ \tilde{q}_{ih} \cdot (1 - p_{ih}^+ - p_{ih}^-), & h \in E_c, h \neq i \\ \lambda_i^\pm \alpha_k^{(ii)\pm}, & h = (i, i, k, \pm) \\ \tilde{q}_{ij} \cdot p_{ij}^\pm \cdot \alpha_k^{(ij)\pm}, & h = (i, j, k, \pm) \end{cases} \quad (4)$$

for $i \in E_c$ as well as

$$q_{(i,j,k,\pm),(i,j,l,\pm)} = T_{kl}^{(ij)\pm} \quad \text{and} \quad q_{(i,j,k,\pm),j} = \eta_k^{(ij)\pm} \quad (5)$$

for $i, j \in E_c$ and $1 \leq k, l \leq m_{ij}^\pm$. For later use we define $q_i := -q_{ii}$ for all $i \in E$.

Denote the MAP constructed in such a way by $(\mathcal{X}, \mathcal{J})$. The original level process \tilde{X} is retrieved via the time change

$$c(t) := \int_0^t 1_{J_s \in E_c} ds, \quad c^{-1}(s) := \inf\{t \geq 0 : c(t) > s\} \quad \text{and} \quad \tilde{X}_t = X_{c^{-1}(t)} \quad (6)$$

for all $t \geq 0$. The inverses of the cumulant functions ψ_i can be given explicitly as

$$\phi_i(\beta) = \begin{cases} \pm\beta, & i \in E_\pm \\ \frac{\beta}{\mu_i}, & i \in E_p \cup E_n \\ \frac{1}{\sigma_i} \sqrt{2\beta + \frac{\mu_i^2}{\sigma_i^2}} - \frac{\mu_i}{\sigma_i}, & i \in E_\sigma \end{cases} \quad (7)$$

We shall, however, use them only for the so-called ascending phases $i \in E_a := E_+ \cup E_p \cup E_\sigma$, cf. [10], chapter VII.

Example 1. Regarding the sample paths in figures 1 and 2, we find \mathcal{J} in phases $i \in E_+$ during the intervals [100, 115], [355, 363], and [523, 538]. In all other intervals, \mathcal{J} is in phases $i \in E_c$. Regarding the time change, we observe in figure 1 a first jump of height 15 at time $t_1 = 100$. We thus obtain $c^{-1}(t_1) = 115$ and $\tilde{X}_{100} = X_{115}$.

2.2. First passage times

Of central use in the present paper will be the recent derivation of the Laplace transforms for the first passage times of MAPs with phase-type jumps as given in [11]. We call the phases $i \in E_d := E_n \cup E_-$ descending. Define $\tilde{\tau}(x) := \inf\{t \geq 0 : \tilde{X}_t > x\}$ for all $x \geq 0$ and assume that $\tilde{X}_0 = 0$. Note that this is the first passage time over the level x for the original MAP \tilde{X} , meaning that we do not count the time spent in jump phases $i \in E_\pm$. This means that $\tilde{\tau}(x) = c(\tau(x)) = \int_0^{\tau(x)} 1_{J_s \in E_c} ds$, according to (6). For $\gamma \geq 0$ denote

$$\mathbb{E}_{ij}(e^{-\gamma\tilde{\tau}(x)}) := \mathbb{E}(e^{-\gamma\tilde{\tau}(x)}; J_{\tau(x)} = j | J_0 = i, X_0 = 0)$$

for all $i, j \in E$. Let $\mathbb{E}(e^{-\gamma\tilde{\tau}(x)})$ denote the matrix with these entries and write

$$\mathbb{E}(e^{-\gamma\tilde{\tau}(x)}) = \begin{pmatrix} \mathbb{E}_{(a,a)}(e^{-\gamma\tilde{\tau}(x)}) & \mathbb{E}_{(a,d)}(e^{-\gamma\tilde{\tau}(x)}) \\ \mathbb{E}_{(d,a)}(e^{-\gamma\tilde{\tau}(x)}) & \mathbb{E}_{(d,d)}(e^{-\gamma\tilde{\tau}(x)}) \end{pmatrix}$$

in obvious block notation with respect to the subspaces $E_a = E_+ \cup E_p \cup E_\sigma$ (ascending phases) and $E_d = E_n \cup E_-$ (descending phases).

Since a first passage to a level above cannot occur in a descending phase, we obtain first $\mathbb{P}(J_{\tau(x)} = j) = 0$ for all $j \in E_d$ and thus $\mathbb{E}_{(d,d)}(e^{-\gamma\tilde{\tau}(x)}) = \mathbb{E}_{(a,d)}(e^{-\gamma\tilde{\tau}(x)}) = \mathbf{0}$ where $\mathbf{0}$ denotes a zero matrix of suitable dimension. Equation (6) in [11] states that

$$\mathbb{E}_{(d,a)}(e^{-\gamma\tilde{\tau}(x)}) = A(\gamma)e^{U(\gamma)x} \quad \text{and} \quad \mathbb{E}_{(a,a)}(e^{-\gamma\tilde{\tau}(x)}) = e^{U(\gamma)x} \quad (8)$$

for some sub-generator matrix $U(\gamma)$ of dimension $|E_a| \times |E_a|$ and a sub-transition matrix $A(\gamma)$ of dimension $|E_d| \times |E_a|$, where $|S|$ denotes the number of elements of a set S . Altogether we can write

$$\mathbb{E}(e^{-\gamma\tilde{\tau}(x)}) = \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} \begin{pmatrix} e^{U(\gamma)x} & \mathbf{0} \end{pmatrix} \quad (9)$$

where I_a denotes the identity matrix of dimension $|E_a| \times |E_a|$.

Write $\Delta_q := \text{diag}(q_i)_{i \in E}$ and let $P = \Delta_q^{-1}Q + I$ denote the transition matrix of phase changes. Note that $p_{ii} = 0$ for all $i \in E$. Let e'_i denote the i th canonical row base vector, with appropriate dimension according to context. According to theorem 3 in [11], $A(\gamma)$ and $U(\gamma)$ satisfy the following equations:

$$\begin{aligned} e'_h U(\gamma) &= \sum_{l=1}^{m_{ij}^+} T_{kl}^{(ij)^+} e'_{(i,j,l,+)} + \eta_k^{(ij)^+} e'_j \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} \quad \text{for } h = (i, j, k, +) \in E_+, \\ e'_i U(\gamma) &= -\phi_i(q_i + \gamma)e'_i + \phi_i(q_i) \sum_{j \in E} p_{ij} e'_j \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} L_i(-U(\gamma)) \quad \text{for } i \in E_p \cup E_\sigma, \\ e'_i A(\gamma) &= \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} ((q_i + \gamma)I - \psi_i(-U(\gamma)))^{-1} \quad \text{for } i \in E_n, \\ e'_i A(\gamma) &= \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} (q_i I - \psi_i(-U(\gamma)))^{-1} \quad \text{for } i \in E_-. \end{aligned}$$

For the MAP $(\mathcal{X}, \mathcal{J})$ with continuous level process, the matrix function

$$L_i(-U(\gamma)) = \frac{q_i}{\phi_i(q_i)} \cdot (\phi_i(q_i + \gamma)I + U(\gamma)) \cdot ((q_i + \gamma)I - \psi_i(-U(\gamma)))^{-1}$$

can be simplified considerably. For $i \in E_\sigma$, the same arguments as in [11], example 2, lead to

$$L_i(-U(\gamma)) = \phi_i^*(q_i) \cdot (\phi_i^*(q_i + \gamma)I - U(\gamma))^{-1} \quad (10)$$

with

$$\phi_i^*(\beta) = \frac{1}{\sigma_i} \sqrt{2\beta + \frac{\mu_i^2}{\sigma_i^2} + \frac{\mu_i}{\sigma_i^2}} \quad (11)$$

Furthermore, $L_i(-U(\gamma)) = I$ for $i \in E_p$ (see example 3 in [11]), while according to (3) $\psi_i(-U(\gamma)) = -\mu_i U(\gamma)$ for $i \in E_n$, and $\psi_i(-U(\gamma)) = U(\gamma)$ for $i \in E_-$. Hence the equations above involve rather simple expressions only.

Considering (7), it is shown in theorem 2 of [11] that the matrices $A(\gamma)$ and $U(\gamma)$ can be determined by successive approximation as the limit of the sequence $((A_n, U_n) : n \geq 0)$ with initial values $A_0 := \mathbf{0}$, $U_0 := -diag(\phi_i(q_i + \gamma)\mathbf{1}_{i \in E_\sigma \cup E_p} + \phi_i(q_i)\mathbf{1}_{i \in E_+})_{i \in E_\sigma}$ and the following iteration:

$$e'_h U_{n+1} = \sum_{l=1}^{m_{ij}^+} T_{kl}^{(ij)+} e'_{(i,j,l,+)} + \eta_k^{(ij)+} e'_j \begin{pmatrix} I_a \\ A_n \end{pmatrix} \quad \text{for } h = (i, j, k, +) \in E_+,$$

$$e'_i U_{n+1} = -\frac{q_i + \gamma}{\mu_i} e'_i + \frac{1}{\mu_i} \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A_n \end{pmatrix} \quad \text{for } i \in E_p,$$

$$e'_i A_{n+1} = \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A_n \end{pmatrix} ((q_i + \gamma)I + \mu_i U_n)^{-1} \quad \text{for } i \in E_n,$$

$$e'_i A_{n+1} = \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A_n \end{pmatrix} (q_i I - U_n)^{-1} \quad \text{for } i \in E_-, \text{ and}$$

$$e'_i U_{n+1} = -\phi_i(q_i + \gamma) e'_i + \frac{2}{\sigma_i^2} \sum_{j \in E, j \neq i} q_{ij} e'_j \begin{pmatrix} I_a \\ A_n \end{pmatrix} (\phi_i^*(q_i + \gamma)I - U_n)^{-1}$$

for $i \in E_\sigma$. For the last equality the relation $\phi_i(q_i)\phi_i^*(q_i) = 2q_i/\sigma_i^2$ has been used. Note that the only difference between the iterations for E_n and E_- is the missing γ in the last factor for E_- , reflecting that we do not discount the time for phases $i \in E_-$ as they are jump phases in real time.

2.3. Time-reversed MAPs

Denote the number of phases in E by $m := |E|$. Let $\pi = (\pi_1, \dots, \pi_m)$ denote the stationary phase distribution, which can be computed by $\pi Q = \mathbf{0}$ and $\pi \mathbf{1} = \sum_{i=1}^m \pi_i = 1$, where $\mathbf{0}$ denotes the zero row vector and $\mathbf{1}$ the column vector with all entries being one. Define the matrix $Q^* = (q_{ij}^*)_{i,j \in E}$ by $q_{ij}^* := \pi_j q_{ji} / \pi_i$ for all $i, j \in E$ or in shorter notation $Q^* := \Delta_\pi^{-1} Q' \Delta_\pi$, where $\Delta_\pi = diag(\pi_1, \dots, \pi_m)$ is the diagonal matrix with entry π_i in its

i th row and the superscript $'$ denotes transposition of a matrix. Then the Markov process with state space E and generator matrix Q^* is a time-reversed version of the original phase process \mathcal{J} . We denote it by $\mathcal{J}^* = (J_t^* : t \geq 0)$.

Based on \mathcal{J}^* we define a time-reversal $(\mathcal{X}^*, \mathcal{J}^*)$ of the original MAP $(\mathcal{X}, \mathcal{J})$ by the rule that \mathcal{X}^* evolves like a Lévy process with parameters $-\mu_i$ (drift) and σ_i^2 (variation) during intervals when the time-reversed phase J_t^* equals $i \in E$. Note that the sign change of the μ_i leads to $E_{\pm}^* = E_{\mp}$, $E_p^* = E_n$, $E_n^* = E_p$, and $E_{\sigma}^* = E_{\sigma}$. We denote the first passage times for $(\mathcal{X}^*, \mathcal{J}^*)$ by $\tau^*(x) := \inf\{t \geq 0 : X_t^* > x\}$ for any level $x \geq 0$.

The same arguments as for equation (3.3) in [2] yield the following relation between the occupation measure (before $\tau(0)$) for the MAP $(\mathcal{X}, \mathcal{J})$ and the first passage time for its time-reversal $(\mathcal{X}^*, \mathcal{J}^*)$:

$$\begin{aligned} \pi_j \mathbb{P}(X_t^* \in dx, X_t^* > X_u^* \forall u < t, J_t^* = i | X_0^* = 0, J_0^* = j) \\ = \pi_i \mathbb{P}(X_t \in -dx, \tau(0) > t, J_t = j | X_0 = 0, J_0 = i) \end{aligned}$$

Write \mathbb{E}_i for the conditional expectation given $X_0 = 0, J_0 = i$. Multiplying by $e^{-\gamma t}$ and integrating over t yields

$$\begin{aligned} \pi_i \mathbb{E}_i \int_0^{\tau(0)} 1_{X_t \in -dx, J_t = j} e^{-\gamma t} dt &= \pi_j \mathbb{E}_j \int_0^{\infty} 1_{X_t^* \in dx, X_t^* > X_u^* \forall u < t, J_t^* = i} e^{-\gamma t} dt \\ &= \pi_j \mathbb{E}_j \left(e^{-\gamma \tau^*(x)}; J_t^* = i \right) \end{aligned} \quad (12)$$

A well-known result that we shall use in the next section is the following lemma which is theorem VI.5(i) and theorem VII.4(i) of [10] applied to Brownian motion and its time-reversal. This lemma also yields an alternative explanation for (10).

Lemma 1. *Let $\mathcal{B} = (B_t : t \geq 0)$ denote a Brownian motion with drift $\mu \in \mathbb{R}$ and variation $\sigma^2 > 0$. Assume that $B_0 = 0$. Further let $\mathcal{E}(q)$ denote a random variable which is independent of \mathcal{B} and has an exponential distribution with parameter $q > 0$. Write*

$$\bar{B}_{\mathcal{E}(q)} := \max_{0 \leq t \leq \mathcal{E}(q)} B_t \quad \text{and} \quad \bar{B}_{\mathcal{E}(q)}^* := \bar{B}_{\mathcal{E}(q)} - B_{\mathcal{E}(q)}$$

as well as

$$G_{\mathcal{E}(q)} := \sup\{t < \mathcal{E}(q) : B_t = \bar{B}_{\mathcal{E}(q)}\} \quad \text{and} \quad G_{\mathcal{E}(q)}^* := \mathcal{E}(q) - G_{\mathcal{E}(q)}$$

Then the pairs $(G_{\mathcal{E}(q)}, \bar{B}_{\mathcal{E}(q)})$ and $(G_{\mathcal{E}(q)}^*, \bar{B}_{\mathcal{E}(q)}^*)$ are independent with respective measures

$$\mathbb{E} \left(e^{-\gamma G_{\mathcal{E}(q)}}; \bar{B}_{\mathcal{E}(q)} \in dx \right) = \phi(q) e^{-\phi(q+\gamma)x} dx$$

and

$$\mathbb{E} \left(e^{-\gamma G_{\mathcal{E}(q)}^*}; \bar{B}_{\mathcal{E}(q)}^* \in dy \right) = \phi^*(q) e^{-\phi^*(q+\gamma)y} dy$$

for $\gamma \geq 0$, where $\phi(\beta)$ and $\phi^*(\beta)$ for $\beta > 0$ are given in the last line of (7) and in (11), respectively.

3. Main result

Let $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}}) = ((\tilde{X}_t, \tilde{J}_t) : t \geq 0)$ denote a MAP with phase-type jumps and assume that $\tilde{X}_0 = 0$. Denote the phase space of $\tilde{\mathcal{J}}$ by \tilde{E} and its generator matrix by \tilde{Q} . Let

$$\tilde{\tau}(u) := \inf\{t > 0 : \tilde{X}_t > u\}$$

denote the first passage time over some level $u \geq 0$. Write

$$\tilde{M}_{\tilde{\tau}(u)} := \sup\{\tilde{X}_t : t < \tilde{\tau}(u)\}$$

for the maximum of $\tilde{\mathcal{X}}$ before the first passage over u . We necessarily have $0 \leq \tilde{M}_{\tilde{\tau}(u)} \leq u$, where $\tilde{M}_{\tilde{\tau}(u)} = 0$ means that $\tilde{\mathcal{X}}$ does not exceed its initial value 0 before it jumps (from a negative value) over the threshold $u \geq 0$. The case $\tilde{M}_{\tilde{\tau}(u)} = u$ means that passage occurs by creeping, i.e. the threshold u is not passed by a jump but continuously. We are further interested in

$$\tilde{G}_{\tilde{\tau}(u)} := \sup\{t < \tilde{\tau}(u) : \tilde{X}_t = \tilde{M}_{\tilde{\tau}(u)}\}$$

which is the time of attaining the maximum before passage over u (cf. lemma 2 regarding its uniqueness). Finally, we wish to determine the density function of the undershoot and the overshoot, defined as

$$U_{\tilde{\tau}(u)} := u - \tilde{X}_{\tilde{\tau}(u)-} \quad \text{and} \quad O_{\tilde{\tau}(u)} := \tilde{X}_{\tilde{\tau}(u)} - u$$

respectively. Our aim is to derive a computable expression for the joint law of these five variables in terms of the measure

$$\mathbb{E} \left(e^{-\gamma \tilde{G}_{\tilde{\tau}(u)} - \gamma^*(\tilde{\tau}(u) - \tilde{G}_{\tilde{\tau}(u)})}; \tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy \right)$$

where $\gamma, \gamma^* \geq 0$ are the arguments for the double Laplace transform, $x, y \geq 0$, and $0 \leq z \leq u$. Note that necessarily $x \geq u - z$.

The approach of analysis in this paper is the same as in [13] for Lévy processes. We divide the sample paths into three parts: the path until the supremum $\tilde{M}_{\tilde{\tau}(u)} < u$ is attained, the path from $\tilde{M}_{\tilde{\tau}(u)}$ to $\tilde{X}_{\tilde{\tau}(u)-}$, and the final jump which leads to an overshoot of the level u .

Remark 1. This decomposes the sample paths either at points of phase changes or at a maximum within the (exponentially distributed) length of a phase regime. Given the phase process, the parts between phase changes are independent according to (1). The standard result in lemma 1 shows that further the parts before and after a maximum within a single phase regime are independent. Thus we obtain conditional independence of all three parts in our path decomposition given the phase process. For more results on this, see [18].

Example 2. Looking at figure 2, the path is decomposed into the parts over the time intervals $[0, 371]$, $[371, 523]$, and $[523, 538]$. Assume that the path is derived from a MAP on a phase space E with 2 phases, one in E_+ and one in E_σ . Then the different phase regimes cover the intervals $[0, 100]$, $[100, 115]$, $[115, 355]$, $[355, 363]$, $[363, 523]$, and $[523, 538]$. These are conditionally independent given the phase process. Further, within the regime $i \in E_\sigma$ over the interval $[363, 523]$, the part $[363, 371]$ until the maximum is independent of the part $[371, 523]$ after the maximum, see lemma 1. As a consequence, the parts $[0, 371]$, $[371, 523]$, and $[523, 538]$ are conditionally independent given the phase process.

The first part of the decomposition is simply a first passage problem. The second part can be determined by (4) via the time-reversed process. Between the three parts we need to take possible (and necessary) phase changes into account. This reasoning will be similar to [9]. A difference to [13] is that we measure the times $\tilde{G}_{\tilde{\tau}(u)}$ and $\tilde{\tau}(u) - \tilde{G}_{\tilde{\tau}(u)}$ in terms of their Laplace transforms. This enables us to provide an explicit formula with expressions that can be readily computed.

One preliminary lemma that we need for the main result concerns the time of attaining the maximum $\tilde{M}_{\tilde{\tau}(u)}$. See [26], lemma 2, for the equivalent statement regarding Lévy processes.

Lemma 2. Define $\tilde{G}'_{\tilde{\tau}(u)} := \inf\{t < \tilde{\tau}(u) : \tilde{X}_t = \tilde{M}_{\tilde{\tau}(u)}\}$. Then $\tilde{G}_{\tilde{\tau}(u)} = \tilde{G}'_{\tilde{\tau}(u)}$ almost surely.

Proof. Since all possible jumps are phase-type and the zero process as well as compound Poisson processes are excluded under any regime, the transition probabilities between levels of local maxima of \tilde{X} are absolutely continuous. Thus the probability of attaining the same

local maximum level twice is 0.

Given the MAP $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$, we construct the modified MAP $(\mathcal{X}, \mathcal{J})$ from it as described in section 2.1. Recall that $P = \Delta_q^{-1}Q + I$, where Q denotes the generator matrix of \mathcal{J} , see (4) and (5), and Δ_q is the diagonal matrix with entries $q_i = -q_{ii}$ for all $i \in E$. Define $p_{ij}^{(+,-)} := \delta_{ij}$ for $i \in E_\sigma$ and $p_{ij}^{(+,-)} := p_{ij}$ for $i \in E_+ \cup E_p$, $j \in E_n \cup E_-$. Further define

$$P^{(+,-)} := \left(p_{ij}^{(+,-)} \right)_{i \in E_a, j \in E_\sigma \cup E_d} \quad \text{and} \quad P^{(c,+)} := \left(p_{ij} 1_{i \in E_c} \right)_{i \in E, j \in E_+}$$

The matrices $P^{(+,-)}$ and $P^{(c,+)}$ subsume the transition probabilities from ascending to descending phases and from continuous to positive jump phases, respectively.

Write $\Delta_\phi := \text{diag}(\phi_i(q_i))_{i \in E_a}$ and $\Delta_{\phi^*} := \text{diag}(\phi_i(q_i) 1_{i \in E_p} + \phi_i^*(q_i) 1_{i \in E_\sigma \cup E_n})_{i \in E}$, where $\phi_i^*(q_i) := q_i / (-\mu_i)$ for $i \in E_n$ and $\phi_i^*(q_i)$ is defined in (11) for $i \in E_\sigma$. Further define the block diagonal matrix $T = \text{diag}(T^{(i,j)})_{(i,j) \in E_c \times E_c}$ as well as the column vector $\eta = ((\eta^{(i,j)})^T)_{(i,j) \in E_c \times E_c}^T$. Here $E_c \times E_c$ must be ordered in some way, say lexicographically. Note that this order must be inherited from the order on E_+ . Finally, define the diagonal matrices $\Pi_a^* = \text{diag}(1/\pi_i)_{i \in E_\sigma \cup E_d}$ and $\Pi_c = \text{diag}(\pi_j 1_{j \in E_c})_{j \in E}$. Now we can state the main result:

Theorem 1. *Let $\tilde{\alpha}$ denote the initial phase distribution of a MAP $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$ with phase-type jumps. Define the row vector $\alpha = (\alpha_i : i \in E)$ on the phase space E of the enlarged MAP $(\mathcal{X}, \mathcal{J})$ by $\alpha_i := \tilde{\alpha}_i$ for all $i \in E_c = \tilde{E}$ and $\alpha_i := 0$ for $i \in E_+ \cup E_-$. Then*

$$\begin{aligned} & \mathbb{E} \left(e^{-\gamma \tilde{G}_{\tilde{\tau}(u)} - \gamma^* (\tilde{\tau}(u) - \tilde{G}_{\tilde{\tau}(u)})}; \tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy \right) \\ &= \alpha \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} e^{U(\gamma)z} \Delta_\phi P^{(+,-)} \\ & \quad \left(\Pi_c \begin{pmatrix} A^*(\gamma^*) \\ I_{E_\sigma \cup E_d} \end{pmatrix} e^{U^*(\gamma^*) \cdot (z - (u-x)) \Pi_a^*} \right)' \Delta_{\phi^*} P^{(c,+)} e^{T \cdot (x+y)} \eta \, dx \, dy \, dz \end{aligned}$$

for all $\gamma, \gamma^* \geq 0$, $0 < z < u$, $x > u - z$, and $y > 0$.

Proof. We consider all possible paths leading to the event

$$\{\tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy\}$$

Since the paths of $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$ can be retrieved from those of $(\mathcal{X}, \mathcal{J})$, we may restrict our attention to $(\mathcal{X}, \mathcal{J})$. We shall, however, speak of jumps whenever we refer to linear movements

governed by phases in $E_- \cup E_+$. Recall the first passage time

$$\tau(u) := \inf\{t \geq 0 : X_t > u\} = \min\{t \geq 0 : X_t = u\}$$

where the last equality holds because of path continuity of \mathcal{X} . Define the times

$$\sigma(u) := \sup\{t < \tau(u) : J_t \notin E_+\} \quad \text{and} \quad \sigma'(u) := \inf\{t > \tau(u) : J_t \notin E_+\}$$

The assumption $z < u$ implies $x \geq u - z > 0$. Further, $y > 0$ almost surely since positive jumps are phase-type and hence absolutely continuous. The time $\sigma(u)$ denotes the instant when the final positive jump, which leads \mathcal{X} over the threshold u , begins. The time $\sigma'(u)$ indicates the instant when this final jump ends and we can measure the overshoot $O_{\tilde{\tau}(u)} = \tilde{X}_{\tilde{\tau}(u)} - u = X_{\sigma'(u)} - u$. We further obtain

$$\tilde{M}_{\tilde{\tau}(u)} = M_{\sigma(u)} := \sup\{X_t : t \leq \sigma(u)\}$$

such that we can write

$$\{\tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy\} = \{M_{\sigma(u)} \in dz, X_{\sigma(u)} \in u - dx, X_{\sigma'(u)} \in u + dy\}$$

We shall employ the following path decomposition. First we consider the path up to its maximum (which is attained at a unique time, due to lemma 2). The second part to consider is the path strictly between the time of maximum and $\sigma(u)$. The last part is the jump. Due to remark 1 the three parts are conditionally independent given the phase process.

The initial phase distribution is denoted by the row vector $\alpha = (\alpha_i : i \in E)$ where $\alpha_i = \mathbb{P}(J_0 = i)$ for all $i \in E$. In order to attain $M_{\sigma(u)} \in dz$, a path must first pass the level z . This happens at the artificial time $\tau(z) := \inf\{t \geq 0 : X_t > z\}$. The Laplace transform of the real time $\tilde{\tau}(z) := \int_0^{\tau(z)} 1_{J_s \in E_c} ds$, restricted to $\{J_{\tau(z)} = i\}$ and with argument $\gamma \geq 0$, is given by

$$\mathbb{E} \left(e^{-\gamma \tilde{\tau}(z)}; J_{\tau(z)} = i | X_0 = 0 \right) = \alpha \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} e^{U(\gamma)z} e_i$$

where e_i denotes the i th canonical base column vector of dimension $|E_a|$.

The further development of the path depends on $J_{\tau(z)}$. There are two cases. In the first case $J_{\tau(z)} = i \in E_+ \cup E_p$. Then we need an instantaneous phase change in order to satisfy $M_{\sigma(u)} \in dz$. This must occur while \mathcal{X} still remains in dz . The infinitesimal rates for this to happen are $\phi_i(q_i) dz = q_i/\mu_i dz$ if $i \in E_p$ and $\phi_i(q_i) dz = q_i dz$ if $i \in E_+$, see

(7). The probabilities for the next phase are then given by $p_{ij}^{(+,-)} = p_{ij}$. In the second case, $J_{\tau(z)} = i \in E_\sigma$, there will be no phase change immediately. In order to satisfy $M_{\sigma(u)} \in dz$, we need to stop the upward ladder process of $\mathcal{X}^{(i)}$ with the infinitesimal rate $\phi_i(q_i) dz$, see lemma 1. The remaining path during phase i will stay below the level z and is independent of the path until $\tau(z)$. Denote the phase after this part by $J_{\tau(z)+} = i$. Then the first part of the path is described by

$$\mathbb{E} \left(e^{-\gamma \tilde{\tau}(z)}; J_{\tau(z)+} = i | X_0 = 0 \right) = \alpha \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} e^{U(\gamma)z} \Delta_\phi P^{(+,-)} e_i dz$$

The second part of the path consists of a movement of \mathcal{X} from the level z at time $\tau(z)+$ to the level $u - x$ at time $\sigma(u)-$ without crossing the level z . Then a subsequent phase change from E_c to E_+ occurs while the level process \mathcal{X} is still in $u - dx$. The first event can be described via the time-reversed MAP $(\mathcal{X}^*, \mathcal{J}^*)$ and relation (12). Denote the matrices governing the first passage times of $(\mathcal{X}^*, \mathcal{J}^*)$ by $A^*(\gamma^*)$ and $U^*(\gamma^*)$ for $\gamma^* \geq 0$. They have dimensions $|E_p \cup E_+| \times |E_\sigma \cup E_d|$ and $|E_\sigma \cup E_d|^2$, respectively. We obtain from (12)

$$\begin{aligned} \mathbb{E} \left(e^{-\gamma^* \cdot (\tilde{\tau}(u) - \tilde{\tau}(z))}; J_{\sigma(u)-} = j | J_{\tau(z)+} = i \right) \\ = \frac{\pi_j}{\pi_i} \mathbb{E} \left(e^{-\gamma^* \cdot \tilde{\tau}^*(z - (u-x))}; J_{\tau^*(z - (u-x))}^* = i | J_0^* = j \right) \end{aligned}$$

where $\tau^*(w) := \inf\{t \geq 0 : X_t^* > w\}$ and $\tilde{\tau}^*(w) = \int_0^{\tau^*(w)} 1_{J_s^* \in E_c} ds$. Here we consider only phases $j \in E_c$, since we wish the process \mathcal{J} to change phase from E_c to E_+ at time $\sigma(u)$, and $i \in E_\sigma \cup E_d = E_a^*$ since these are the only phases in which the time-reversed process \mathcal{X}^* can cross the level $z - (u - x)$ from below. In matrix notation, the Laplace transform of the real time spent between $\tilde{\tau}(z)$ and $\tilde{\tau}(u)$ is thus

$$\mathbb{E} \left(e^{-\gamma^* \cdot (\tilde{\tau}(u) - \tilde{\tau}(z))} \right) = \left(\Pi_c \begin{pmatrix} A^*(\gamma^*) \\ I_{E_\sigma \cup E_n} \end{pmatrix} e^{U^*(\gamma^*) \cdot (z - (u-x))} \Pi_a^* \right)' \quad (13)$$

with $\Pi_a^* = \text{diag}(1/\pi_i)_{i \in E_\sigma \cup E_d}$ and $\Pi_c = \text{diag}(\pi_j 1_{j \in E_c})_{j \in E}$. Now to the subsequent phase change: Denote the phase at the time of moving to the level $u - x$ by $J_{\sigma(u)-} = j \in E_c$. In order to trigger the jump that will overshoot the level u with $U_{\tilde{\tau}(u)} \in dx$, the phase process \mathcal{J} needs to change from j to some $k \in E_+$ while the level process \mathcal{X} is still in $u - dx$. This happens with probabilities $q_j/\mu_j dx \cdot p_{jk}$ for $j \in E_p$, $q_j/(-\mu_j) dx \cdot p_{jk}$ for $j \in E_n$, and $\phi_j^*(q_j) dx \cdot p_{jk}$ for $j \in E_\sigma$. Hence the second part of the path can be subsumed in matrix

notation as

$$\begin{aligned} \mathbb{E} \left(e^{-\gamma^* \cdot (\tilde{\tau}(u) - \tilde{\tau}(z))}; J_{\sigma(u)} = k | J_{\tau(z)_+} = i \right) \\ = e'_i \left(\Pi_c \begin{pmatrix} A^*(\gamma^*) \\ I_{E_\sigma \cup E_n} \end{pmatrix} e^{U^*(\gamma^*) \cdot (z - (u-x)) \Pi_a^*} \right)' \Delta_{\phi^*} P^{(c,+)} e_k dx \end{aligned}$$

The last part of the path is merely the final jump. Given $J_{\sigma(u)} = k \in E_+$, it is independent of the path before. In order to be the final jump, it must be larger than x . Given this, the phase-type assumption on the jumps yields the conditional density function $e'_k e^{T \cdot (x+y)} \eta dy$. This completes the proof.

If the process starts with a negative drift, then the singular case $\tilde{M}_{\tilde{\tau}(u)} = 0$ is possible. This implies $\tilde{G}_{\tilde{\tau}(u)} = 0$ and $x > u$. The remaining triple law is given in the following theorem. For $\gamma^* = 0$ and $E_\sigma = \emptyset$ it yields equation (3.6) in [2] and theorem 1 in [25].

Corollary 1. *Let α be an initial phase distribution with support on E_n . Then*

$$\begin{aligned} \mathbb{E} \left(e^{-\gamma^* \tilde{\tau}(u)}; \tilde{M}_{\tilde{\tau}(u)} = 0, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy \right) \\ = \alpha \left(\Pi_c \begin{pmatrix} A^*(\gamma^*) \\ I_{E_\sigma \cup E_d} \end{pmatrix} e^{U^*(\gamma^*) \cdot (x-u) \Pi_a^*} \right)' \Delta_{\phi^*} P^{(c,+)} e^{T \cdot (x+y)} \eta dx dy \end{aligned}$$

for $x > u$ and $y > 0$.

Two other singular cases that may arise are given in the following corollaries. The reasoning for them is the same as for theorem 1.

Corollary 2. *Passage by creeping*

$$\mathbb{E} \left(e^{-\gamma \tilde{\tau}(u)}; \tilde{M}_{\tilde{\tau}(u)} = u, U_{\tilde{\tau}(u)} = 0, O_{\tilde{\tau}(u)} = 0 \right) = \alpha \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} e^{U(\gamma)u} \mathbf{1}_{E_p \cup E_\sigma}$$

where $\mathbf{1}_{E_p \cup E_\sigma}$ is a column vector of dimension $|E_a|$ with i th entry being 0 for $i \in E_+$ and 1 for $i \in E_p \cup E_\sigma$.

Corollary 3. *Passage by jump from a running maximum*

$$\begin{aligned} \mathbb{E} \left(e^{-\gamma \tilde{\tau}(u)}; \tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in u - dz, O_{\tilde{\tau}(u)} \in dy \right) \\ = \alpha \begin{pmatrix} I_a \\ A(\gamma) \end{pmatrix} e^{U(\gamma)z} \Delta_{\phi} P^{(p,+)} e^{T \cdot (u-z+y)} \eta dy dz \end{aligned}$$

where $P^{(p,+)} = (p_{ij}1_{i \in E_p})_{i \in E_a, j \in E_+}$.

4. Application to insurance risk

Consider a risk reserve process with initial capital $u \geq 0$ and claims occurring like a Markovian point process (MPP).^{*} It is shown in [7] that the class of MPPs is dense within the class of marked point processes. Thus we incur no serious restriction in generality. Denote the claim arrival process by $(\mathcal{N}, \tilde{\mathcal{J}}) = ((N_t, \tilde{\mathcal{J}}_t) : t \geq 0)$ and the phase space for $\tilde{\mathcal{J}}$ by \tilde{E} . Assume that the claim sizes have a phase-type distribution and denote the n th claim size by C_n , $n \in \mathbb{N}$. By [28] the class of phase-type distributions is dense within the class of all distributions on the positive real numbers. We assume further that the premium income between claims can be modelled by a Brownian motion, where the parameters $\tilde{\mu}_i$ (drift) and $\tilde{\sigma}_i$ (variation) at time t may depend on the current phase $\tilde{\mathcal{J}}_t = i$ of the claim arrival process. For insurance risk we typically have $\tilde{\mu}_i > 0$ for all $i \in \tilde{E}$. We shall allow $\tilde{\sigma}_i = 0$ for some (or possibly all) phases, under which condition the Brownian motion becomes a linear drift. However, the case $\tilde{\mu}_i = \tilde{\sigma}_i = 0$ of a constant (null) movement shall be excluded. Then the process of premium income is a Markov-modulated Brownian motion which we denote by $(\mathcal{B}, \tilde{\mathcal{J}}) = ((B_t, \tilde{\mathcal{J}}_t) : t \geq 0)$. We assume that $B_0 = 0$.

Note that $\tilde{\mathcal{J}}$ here is the same as for the claim arrival process $(\mathcal{N}, \tilde{\mathcal{J}})$. This is no restriction in modelling power as we can choose identical parameters $(\tilde{\mu}_i, \tilde{\sigma}_i) = (\tilde{\mu}_j, \tilde{\sigma}_j)$ for different phases $i \neq j \in \tilde{E}$ and map two different environments for the premium income and the claim arrivals by using Kronecker products. Rather on the contrary, a common phase space enables us to model correlations between claim arrivals, claim sizes, and the premium income.

With the definitions above, the risk reserve process $\mathcal{U} = (U_t : t \geq 0)$ is given by

$$U_t = u + B_t - \sum_{n=1}^{N_t} C_n$$

for $t \geq 0$. Denote the net claim process by $\tilde{\mathcal{X}} = (\tilde{X}_t : t \geq 0)$ where $\tilde{X}_t := u - U_t$ for all $t \geq 0$. Then the process $(\tilde{\mathcal{X}}, \tilde{\mathcal{J}})$ is a MAP with phase-type jumps and we can apply the analysis presented in section 3. For $\gamma = \gamma^*$ this yields in particular the Gerber–Shiu function

^{*} see [24, 23, 7]. This has traditionally been called Markovian arrival process and abbreviated as MAP. Since we use the shortcut MAP for the more general class of Markov-additive processes already, we prefer to use the term Markovian point process and the abbreviation MPP instead. Some authors use the abbreviation MArP.

after integrating out the variable $\tilde{M}_{\tilde{\tau}(u)}$. The following examples shall illustrate this.

Example 3. We consider the classical compound Poisson model. Inter-claim times and claim sizes are iid exponential with parameter $\lambda > 0$ and $\beta > 0$, respectively. The rate of premium income is $c > 0$. The net profit condition is then $\lambda/(c\beta) < 1$. This model has been examined in [16]. The net claim amount at time $t \geq 0$ is given by

$$\tilde{X}_t = \sum_{n=0}^{N_t} C_n - ct \quad (14)$$

where $(N_t : t \geq 0)$ is a Poisson process with intensity λ and the $C_n, n \in \mathbb{N}$, are iid random variables with exponential distribution of parameter β .

The process of accumulated claims can be analysed as a MAP with exponential (and hence phase-type) positive jumps, where $T = -\beta$ and $\eta = \beta$. We further obtain the MAP $(\mathcal{X}, \mathcal{J})$ as follows. Let the phase space be given by $E_+ = \{1\}$, $E_n = \{2\}$, and $E_\sigma = \emptyset$. The parameters are given by $\sigma_1 = \sigma_2 = 0$, $\mu_1 = 1$, $\mu_2 = -c$, $\nu_1 = \nu_2 = \mathbf{0}$, and

$$Q = \begin{pmatrix} -\beta & \beta \\ \lambda & -\lambda \end{pmatrix}$$

Note that phase 1 represents the upwards jumps and we will not discount the time during sojourns in it. As shown in [11], example 5, the Laplace transform of the first passage time $\tilde{\tau}(x) := \inf\{t \geq 0 : \tilde{X}_t > x\}$ to a level $x > 0$ is given by

$$\mathbb{E}(e^{-\gamma\tilde{\tau}(x)}) = A(\gamma)e^{U(\gamma)x} \quad \text{where} \quad A(\gamma) = \frac{\beta - R}{\beta}, \quad U(\gamma) = -R$$

and

$$-R = \frac{1}{2c} \left(\lambda + \gamma - c\beta - \sqrt{(c\beta - \gamma - \lambda)^2 + 4c\beta\gamma} \right)$$

which coincides with equation (4.24) in [16], noting that γ is denoted as δ there. The time-reversed process has a positive drift ($E_p^* = E_n = \{2\}$) and negative jumps ($E_-^* = E_+ = \{1\}$). Instead of evaluating the rather complicated expression (13), we can treat it as a spectrally negative Lévy process directly. Its cumulant function is given by

$$\psi^*(x) = cx - \lambda + \lambda \frac{\beta}{\beta + x}$$

Hence we find for the inverse of ψ^*

$$\phi^*(\gamma) = \frac{1}{2c} \left(\lambda + \gamma - c\beta + \sqrt{(c\beta - \lambda - \gamma)^2 + 4c\beta\gamma} \right)$$

which is denoted by ρ in [16], equation (3.12). We thus obtain

$$\left(\Delta_\pi \begin{pmatrix} A^*(\gamma) \\ I_{E_n} \end{pmatrix} e^{U^*(\gamma) \cdot (z - (u-x))} \Delta_\pi^{-1} \right)^T = e^{-\rho \cdot (z - (u-x))} \quad (15)$$

Further we shall need $\phi_1(q_1) = \beta$ and $\phi_2^*(q_2) = \lambda/c$.

In order to find the function $f(x|u) := \mathbb{E}(e^{-\gamma\tilde{\tau}(u)}; U_{\tilde{\tau}(u)} \in dx)$, we consider the cases $x \leq u$ and $x > u$ separately. We always have $\tilde{M}_{\tilde{\tau}(u)} \geq u - x$. From theorem 1 and the above equation (15) we compute the marginal density function

$$\begin{aligned} \mathbb{E}(e^{-\gamma\tilde{\tau}(u)}; U_{\tilde{\tau}(u)} \in dx, x \leq u) &= \int_{u-x}^u A(\gamma) e^{U(\gamma)z} \phi_1(q_1) e^{-\rho \cdot (z - u + x)} \phi_2^*(q_2) dz e^{Tx} \\ &= \int_{u-x}^u \frac{\beta - R}{\beta} e^{-Rz} \beta e^{-\rho \cdot (z - u + x)} \frac{\lambda}{c} dz e^{-\beta x} \\ &= \frac{\lambda}{c} (\beta - R) e^{-\rho \cdot (x - u)} e^{-\beta x} \int_{u-x}^u e^{-(R+\rho)z} dz \\ &= \frac{\lambda}{c} \frac{\beta - R}{R + \rho} e^{-(\rho+\beta)x} e^{\rho u} \left(e^{-(R+\rho) \cdot (u-x)} - e^{-(R+\rho) \cdot u} \right) \\ &= \frac{\lambda}{c} \frac{\beta - R}{R + \rho} e^{-(\rho+\beta)x} \left(e^{(R+\rho) \cdot x} - 1 \right) e^{-Ru} \end{aligned}$$

which coincides with equation (6.40) in [16]. Now for the case $x > u$. This means that the event $M_{\tilde{\tau}(u)} = 0$ may attain a positive probability. Corollary 1 and equation (15) yield for this case

$$\begin{aligned} \mathbb{E}(e^{-\gamma\tilde{\tau}(u)}, U_{\tilde{\tau}(u)} \in dx, \tilde{M}_{\tilde{\tau}(u)} = 0) &= e^{-\rho \cdot (x-u)} \phi_2^*(q_2) e^{Tx} = e^{-\rho \cdot (x-u)} \frac{\lambda}{c} e^{-\beta x} \\ &= \frac{\lambda}{c} e^{-(\rho+\beta)x} e^{\rho u} \end{aligned}$$

For the case of $x > u$ and $M_{\tilde{\tau}(u)} > 0$, theorem 1 yields

$$\begin{aligned} \mathbb{E}(e^{-\gamma\tilde{\tau}(u)}, U_{\tilde{\tau}(u)} \in dx, \tilde{M}_{\tilde{\tau}(u)} > 0, x > u) &= \int_0^u \frac{\beta - R}{\beta} e^{-Rz} \beta e^{-\rho \cdot (z - u + x)} \frac{\lambda}{c} dz e^{-\beta x} \\ &= \frac{\lambda}{c} \frac{\beta - R}{R + \rho} e^{-(\rho+\beta)x} e^{\rho u} \left(1 - e^{-(R+\rho) \cdot u} \right) \\ &= \frac{\lambda}{c} \frac{\beta - R}{R + \rho} e^{-(\rho+\beta)x} \left(e^{\rho u} - e^{-Ru} \right) \end{aligned}$$

Adding these results we obtain finally

$$\begin{aligned} \mathbb{E}(e^{-\gamma\tilde{\tau}(u)}, U_{\tilde{\tau}(u)} \in dx, x > u) &= \frac{\lambda}{c} e^{-(\rho+\beta)x} \left(e^{\rho u} + \frac{\beta - R}{R + \rho} (e^{\rho u} - e^{-Ru}) \right) \\ &= \frac{\lambda}{c \cdot (R + \rho)} e^{-(\rho+\beta)x} \left((\beta + \rho) e^{\rho u} - (\beta - R) e^{-Ru} \right) \end{aligned}$$

which coincides with (6.39) in [16].

Example 4. The same net claim process as in (14) can be analysed by the approach in [13], example 8. There the time aspect is neglected and thus we set $\gamma = \gamma^* = 0$. Then we obtain $\rho = 0$ and $-R = (\lambda - c\beta)/c = \lambda/c - \beta$ by (3.12) and (4.24) in [16]. This implies $\beta - R = \lambda/c$, and theorem 1 now yields for $0 < z < u$

$$\begin{aligned} \mathbb{P}(\tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy) &= \frac{\beta - R}{\beta} e^{-Rz} \beta e^{-\rho \cdot (z - (u-x))} \frac{\lambda}{c} e^{-\beta(x+y)} \beta \\ &= \frac{\lambda}{c} e^{-(\beta - \lambda/c)z} \frac{\lambda}{c} e^{-\beta \cdot (x+y)} \beta \end{aligned}$$

Corollary 1 yields further for $x > u$

$$\mathbb{P}(\tilde{M}_{\tilde{\tau}(u)} = 0, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy) = e^{-\rho \cdot (x-u)} \frac{\lambda}{c} e^{-\beta \cdot (x+y)} \beta = \frac{\lambda}{c} e^{-\beta \cdot (x+y)} \beta$$

The form obtained in [13], p.101, is

$$\mathbb{P}(\tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy) = \frac{1}{c} \sum_{n \geq 0} \nu^{*n}(dz) \Pi_X(x + dy) dx$$

for $0 \leq z \leq u$, $x \geq u - z$ and $y > 0$, where $\Pi_X(dx) = \lambda e^{-\beta x} \beta dx$ is the Lévy measure, $\nu^{*0} = \delta_0$ (the Dirac measure on 0), and ν^{*n} denotes the n -fold convolution of the measure

$$\nu(dx) = \frac{1}{c} \Pi_X(x, \infty) dx = \frac{\lambda}{c} e^{-\beta x} dx$$

It is immediate that the results coincide for the singular case $M_{\tilde{\tau}(u)} = 0$. In order to show agreement for the case $M_{\tilde{\tau}(u)} > 0$, it suffices to show that

$$\sum_{n \geq 1} \nu^{*n}(dz) = \frac{\lambda}{c} e^{-(\beta - \lambda/c)z} dz \quad (16)$$

holds. Taking Laplace transforms on both sides, we obtain on the left-hand side

$$\int e^{-\alpha z} \sum_{n \geq 1} \nu^{*n}(dz) = \sum_{n \geq 1} \int e^{-\alpha z} \nu^{*n}(dz) = \sum_{n \geq 1} (L_\nu(\alpha))^n = \frac{L_\nu(\alpha)}{1 - L_\nu(\alpha)}$$

where $L_\nu(\alpha) = \lambda/c (\alpha + \beta)^{-1}$, $\alpha > 0$, is the Laplace transform of ν . On the right-hand side we obtain

$$\frac{\lambda}{c} \int e^{-\alpha z} e^{-(\beta - \lambda/c)z} dz = \frac{\lambda}{c} \cdot \frac{1}{\alpha + \beta - \frac{\lambda}{c}} = \frac{\frac{\lambda}{c} \cdot \frac{1}{\alpha + \beta}}{1 - \frac{\lambda}{c} \cdot \frac{1}{\alpha + \beta}}$$

which yields (16).

Example 5. The joint density function of undershoot and overshoot has been derived in [9] for the fluid flow case from an insurance perspective. We set $\tilde{X}_t := u - U_t$ where u is the

initial capital and U_t is the risk reserve process as given in [9], p.434. For the resulting MAP $(\mathcal{X}, \mathcal{J})$ constructed from $\tilde{\mathcal{X}}$ as in section 2.1, we obtain $E_p = E_\sigma = E_- = \emptyset$ and $E_+ = S_2$, $E_n = S_1$ where S_i are the notations in [9], p.434. Comparing the notations for the generator matrix of the phase process \mathcal{J} , we get the block partition

$$Q = \begin{pmatrix} Q_{++} & Q_{+-} \\ Q_{-+} & Q_{--} \end{pmatrix} = \begin{pmatrix} T_{22} & T_{21} \\ T_{12} & T_{11} \end{pmatrix} = T$$

Write $P = \Delta_q^{-1}Q + I$ and P_{+-} etc. for the obvious block partitions. The assumption $c = 1$ therein translates to $\mu_i = -1$ for all $i \in E_n$ in our notation. Since we do not need to determine the time variables, we can set $\gamma = \gamma^* = 0$.

First of all we observe that Γ in [9], p.436, equals $A(0)$ by definition. Then equation (16) in [11] yields

$$U(0) = Q_{++} + Q_{+-}A(0) = T_{22} + T_{21}\Gamma = H$$

which is the notation in [9], p.437. Furthermore, $\phi(q_i) = q_i$ for all $i \in E_+$ and $P^{(+,-)} = P_{+-}$, which yields $\Delta_\phi P^{(+,-)} = Q_{+-} = T_{21}$. Regarding the time reversal $(\mathcal{X}^*, \mathcal{J}^*)$ of $(\mathcal{X}, \mathcal{J})$, the phase subspaces translate as $E_p^* = E_n$ and $E_-^* = E_+$. Write

$$Q^* = \Delta_\pi^{-1}Q^T \Delta_\pi = \begin{pmatrix} Q_{--}^* & Q_{-+}^* \\ Q_{+-}^* & Q_{++}^* \end{pmatrix}$$

such that $Q_{++}^* = \Delta_{\pi_-}^{-1}Q_{--}^T \Delta_{\pi_-}$ and $Q_{+-}^* = \Delta_{\pi_-}^{-1}Q_{-+}^T \Delta_{\pi_+}$, denoting $\Delta_{\pi_-} = \text{diag}(\pi_i)_{i \in E_n}$ and $\Delta_{\pi_+} = \text{diag}(\pi_i)_{i \in E_+}$. Then $U^*(0) = Q_{++}^* + Q_{+-}^*A^*(0)$, where $A^*(0)$ has dimension $E_+ \times E_n$. Since $E_\sigma = E_\sigma^* = \emptyset$, we obtain $A^*(0) = \Delta_{\pi_+}^{-1}A^T(0)\Delta_{\pi_-}$. Thus

$$U^*(0) = \Delta_{\pi_-}^{-1} (Q_{--}^T + Q_{-+}^T A^T(0)) \Delta_{\pi_-}$$

and

$$\begin{aligned} \left(\Delta_{\pi_-} e^{U^*(0) \cdot (z - (u-x))} \Delta_{\pi_-}^{-1} \right)^T &= \left(e^{(Q_{--}^T + Q_{-+}^T A^T(0)) \cdot (z - (u-x))} \right)^T \\ &= e^{(Q_{--} + A(0)Q_{+-}) \cdot (z - (u-x))} \\ &= e^{(T_{11} + \Gamma T_{21}) \cdot (z - (u-x))} = e^{K \cdot (z - (u-x))} \end{aligned} \quad (17)$$

in the notation of [9], p.436. Finally, $\phi^*(q_i) = q_i$ for all $E_p^* = E_n$ and $P^{(c,+)} = P_{-+}$ such that $\Delta_{\phi^*} P^{(c,+)} = Q_{-+} = T_{12}$. For the jump the notations translate as $T = Q_{++} = T_{22}$ and $\eta = t_2$.

Hence the density function $h(u, x, y)$ of the surplus prior to ruin (the undershoot x) and the deficit at ruin (the overshoot y) is given by

$$h(u, x, y) = \int_{z=u-x}^u \mathbb{P}(M_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy)$$

for $x < u$. Theorem 1 yields

$$\begin{aligned} h(u, x, y) &= \alpha A(0) \int_{z=u-x}^u e^{U(0)z} \Delta_{\phi} P^{(+,-)} \left(\Delta_{\pi_-} e^{U^*(\gamma^*) \cdot (z-(u-x))} \Delta_{\pi_-}^{-1} \right)^T dz \\ &\quad \Delta_{\phi^*} P^{(c,+)} e^{T \cdot (x+y)} \eta \\ &= \alpha \Gamma e^{H \cdot (u-x)} \int_{w=0}^x e^{Hw} T_{21} e^{Kw} dw T_{12} e^{T_{22} \cdot (x+y)} t_2 \end{aligned}$$

after a substitution $w = z - (u - x)$. This coincides with equation (17) in [9], noting the definition (19) therein. The case $x > u$ includes the singular event $\{\tilde{M}_{\tilde{\tau}(u)} = 0\}$. For this corollary 1 and (17) yield

$$\begin{aligned} \mathbb{P}(\tilde{M}_{\tilde{\tau}(u)} = 0, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy) &= \alpha \left(\Delta_{\pi_-} e^{U^*(\gamma^*) \cdot (z-(u-x))} \Delta_{\pi_-}^{-1} \right)^T \\ &\quad \Delta_{\phi^*} P^{(c,+)} e^{T \cdot (x+y)} \eta \\ &= \alpha e^{K \cdot (x-u)} T_{12} e^{T_{22} \cdot (x+y)} t_2 \end{aligned} \quad (18)$$

The other part can be determined via theorem 1 as

$$\begin{aligned} \mathbb{P}(\tilde{M}_{\tilde{\tau}(u)} = 0, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy) &= \alpha \Gamma \int_{z=0}^u e^{Hz} T_{21} e^{K \cdot (z-(u-x))} dz T_{12} e^{T_{22} \cdot (x+y)} t_2 \\ &= \alpha \Gamma R(u) e^{K \cdot (x-u)} T_{12} e^{T_{22} \cdot (x+y)} t_2 \end{aligned} \quad (19)$$

where $R(u) = \int_{w=0}^u e^{Hw} T_{21} e^{Kw} dw$ is the notation in [9], equation (19). Adding the two results (18) and (19) yields the same expression as (18) in [9].

Example 6. We consider the perturbed version of the classical compound Poisson model with exponential claim sizes. The net claim process is a Lévy process with parameters $\tilde{\sigma} > 0$, $\tilde{\mu} = -c < 0$, and $\tilde{\nu}(dx) = 1_{x>0} \lambda \cdot \beta e^{-\beta x} dx$. From this we construct the MAP $(\mathcal{X}, \mathcal{J})$ according to section 2.1. This has phase space $E = \{0, 1\}$ with $E_+ = \{0\}$ and $E_{\sigma} = \{1\}$. The generator matrix of \mathcal{J} is given by

$$Q = \begin{pmatrix} -\beta & \beta \\ \lambda & -\lambda \end{pmatrix}$$

which implies $q_0 = \beta$ and $q_1 = \lambda$. The parameters for \mathcal{X} are $\sigma_0 = 0$, $\mu_0 = 1$, $\sigma_1 = \tilde{\sigma} =: \sigma$, $\mu_1 = -c$, and $\nu_0 = \nu_1 = \mathbf{0}$. In order to determine the matrix $U = U(\gamma)$, we note that $E_a = E$ such that U has dimension 2×2 and there is no matrix $A(\gamma)$. We need the values of $\phi_1(q_1 + \gamma)$ and $\phi_1^*(q_1 + \gamma)$ which are given in (7) and (11) as

$$\phi_1(q_1 + \gamma) = \frac{1}{\sigma} \sqrt{2(\lambda + \gamma) + \frac{c^2}{\sigma^2}} + \frac{c}{\sigma^2} \quad \text{and} \quad \phi_1^*(q_1 + \gamma) = \frac{1}{\sigma} \sqrt{2(\lambda + \gamma) + \frac{c^2}{\sigma^2}} - \frac{c}{\sigma^2}$$

Then the matrix U is determined by the fixed point equation

$$\begin{aligned} e'_0 U &= -\beta e'_0 + \beta e'_1 = (-\beta, \beta) \\ e'_1 U &= -\phi_1(q_1 + \gamma) e'_1 + \frac{2}{\sigma^2} \lambda e'_0 (\phi_1^*(q_1 + \gamma) I - U)^{-1} \end{aligned} \quad (20)$$

We further obtain

$$\Delta_\phi P^{(+,-)} = \begin{pmatrix} \beta & 0 \\ 0 & \phi_1(q_1) \end{pmatrix} \begin{pmatrix} 0 \\ 1 \end{pmatrix} = \begin{pmatrix} 0 \\ \phi_1(q_1) \end{pmatrix}$$

For the second part in the formula of theorem 1 we need to determine the time-reversed MAP $(\mathcal{X}^*, \mathcal{J}^*)$ according to section 2.3. We obtain $E_\sigma^* = E_\sigma = \{1\}$ and $E_-^* = E_+ = \{0\}$ with parameters $\sigma_0^* = 0$, $\mu_0^* = -1$, $\sigma_1^* = \sigma$, $\mu_1^* = c$ for \mathcal{X}^* . The generator matrix for \mathcal{J}^* turns out to be $Q^* = \Delta_\pi^{-1} Q' \Delta_\pi = Q$, where $\pi = (\lambda/(\beta + \lambda), \beta/(\beta + \lambda))$. Thus $q_0^* = q_0 = \beta$ and $q_1^* = q_1 = \lambda$. Since $E_a^* = \{1\}$ and $E_d^* = \{0\}$, the matrices $U^* = U^*(\gamma^*)$ and $A^* = A^*(\gamma^*)$ are real numbers. They are determined by

$$U^* = -\phi_1^*(\lambda + \gamma^*) + \frac{2}{\sigma^2} \lambda \frac{A^*}{\phi_1(\lambda + \gamma^*) - U^*} \quad \text{and} \quad A^* = \frac{\beta}{\beta - U^*} \quad (21)$$

where the duality of \mathcal{X} and \mathcal{X}^* yields $(\phi_1^*)^* = \phi_1$. With the definitions of

$$\Pi_c = \begin{pmatrix} 0 & 0 \\ 0 & \frac{\beta}{\beta + \lambda} \end{pmatrix}, \quad \Pi_a^* = \frac{\beta + \lambda}{\beta}, \quad \Delta_{\phi^*} = \begin{pmatrix} 0 & 0 \\ 0 & \phi_1^*(\lambda) \end{pmatrix}, \quad P^{(c,+)} = \begin{pmatrix} 0 \\ 1 \end{pmatrix}$$

this yields

$$\left(\Pi_c \begin{pmatrix} A^* \\ 1 \end{pmatrix} e^{U^* \cdot (z - (u-x)) \Pi_a^*} \right)' \Delta_{\phi^*} P^{(c,+)} = e^{U^* \cdot (z - (u-x))} \phi_1^*(\lambda)$$

Finally the parameters for the final jump part are $T = -\beta$ and $\eta = \beta$. For the natural choice $\alpha = e'_1$ the expression in theorem 1 becomes, after the simplification $\phi_1(\lambda) \cdot \phi_1^*(\lambda) = 2\lambda/\sigma^2$,

$$\begin{aligned} \mathbb{E} \left(e^{-\gamma \tilde{G}_{\tilde{\tau}(u)} - \gamma^* (\tilde{\tau}(u) - \tilde{G}_{\tilde{\tau}(u)})}; \tilde{M}_{\tilde{\tau}(u)} \in dz, U_{\tilde{\tau}(u)} \in dx, O_{\tilde{\tau}(u)} \in dy \right) \\ = \left(e'_1 e^{U(\gamma)z} e_1 \right) \frac{2\lambda}{\sigma^2} e^{U^*(\gamma^*) \cdot (z - (u-x))} \cdot \beta e^{-\beta(x+y)} dx dy dz \end{aligned}$$

The number $U^*(\gamma^*)$ and the matrix $U(\gamma)$ are probably best determined numerically from (20) and (21).

References

- [1] AHN, S. AND BADESCU, A. L. (2007). On the analysis of the Gerber-Shiu discounted penalty function for risk processes with Markovian arrivals. *Insurance: Mathematics and Economics* **41**, 234–249.
- [2] ASMUSSEN, S. (1991). Ladder heights and the Markov-modulated M/G/1 queue. *Stoch. Proc. Appl.* **37**, 313–326.
- [3] ASMUSSEN, S. (1995). Stationary distributions via first passage times. In *Advances in queueing: Theory, methods, and open problems*. ed. J. Dshalalow. CRC Press, Boca Raton pp. 79–102.
- [4] ASMUSSEN, S. (2000). *Ruin probabilities*. Singapur: World Scientific.
- [5] ASMUSSEN, S. (2003). *Applied Probability and Queues*. New York etc.: Springer.
- [6] ASMUSSEN, S., AVRAM, F. AND PISTORIUS, M. (2004). Russian and American put options under exponential phase-type Lévy models. *Stochastic Processes and their Applications* **109**, 79–111.
- [7] ASMUSSEN, S. AND KOOLE, G. (1993). Marked point processes as limits of Markovian arrival streams. *J. Appl. Probab.* **30**, 365–372.
- [8] BADESCU, A., BREUER, L., DA SILVA SOARES, A., LATOUCHE, G., REMICHE, M.-A. AND STANFORD, D. (2005). Risk processes analyzed as fluid queues. *Scandinavian Actuarial Journal* 127–141.
- [9] BADESCU, A., BREUER, L., DREKIC, S., LATOUCHE, G. AND STANFORD, D. (2005). The surplus prior to ruin and the deficit at ruin for a correlated risk process. *Scandinavian Actuarial Journal* 433–445.
- [10] BERTOIN, J. (1996). *Lévy Processes*. Cambridge University Press, Cambridge etc.

- [11] BREUER, L. (2008). First passage times for Markov-additive processes with positive jumps of phase-type. *J. Appl. Prob.* **45**, 779–799.
- [12] CHIU, S. AND YIN, C. (2003). The time of ruin, the surplus prior to ruin and the deficit at ruin for the classical risk process perturbed by diffusion. *Insurance: Mathematics and Economics* **33**, 59–66.
- [13] DONEY, R. AND KYPRIANOU, A. (2006). Overshoots and undershoots of Lévy processes. *Annals of Applied Probability* **16**, 91–106.
- [14] GARRIDO, J. AND MORALES, M. (2006). On the expected discounted penalty function for Lévy risk processes. *North American Actuarial Journal* **10**, 196–216.
- [15] GERBER, H. AND LANDRY, B. (1998). On the discounted penalty at ruin in a jump-diffusion and the perpetual put option. *Insurance: Mathematics and Economics* **22**, 263–276.
- [16] GERBER, H. AND SHIU, E. (1998). On the time value of ruin. *North American Actuarial Journal* **2**, 48–78.
- [17] GERBER, H. AND SHIU, E. (2005). The time value of ruin in a Sparre Andersen model. *North American Actuarial Journal* **9**, 49–84.
- [18] KLUSIK, P. AND PALMOWSKI, Z. A note on Wiener-Hopf factorization for Markov Additive processes. *arXiv* 0906.1223v1.
- [19] KYPRIANOU, A. AND PALMOWSKI, Z. (2008). Fluctuations of spectrally negative Markov additive processes. *Séminaire de Probabilités XLI (Lecture Notes Math.)* **1934**, 121–135.
- [20] LI, S. AND GARRIDO, J. (2005). The Gerber-Shiu function in a Sparre Andersen risk process perturbed by diffusion. *Scand. Act. J.* 161–186.
- [21] LU, Y. AND TSAI, C. C.-L. (2007). The expected discounted penalty at ruin for a Markov-modulated risk process perturbed by diffusion. *North American Actuarial Journal* **11**, 136–152.

- [22] LU, Y., WU, R. AND XU, R. (2006). The joint distributions of some extrema for the classical risk process perturbed by diffusion. *Chinese Journal of Engineering Mathematics* **23**, 355–360.
- [23] LUCANTONI, D. M. (1991). New results on the single server queue with a batch Markovian arrival process. *Commun. Stat., Stochastic Models* **7**, 1–46.
- [24] NEUTS, M. F. (1979). A versatile Markovian point process. *J. Appl. Probab.* **16**, 764–774.
- [25] NG, A. AND YANG, H. (2006). On the joint distribution of surplus before and after ruin under a Markovian regime switching model. *Stoch. Proc. Appl.* **116**, 244–266.
- [26] PECHERSKII, E. AND ROGOZIN, B. (1969). On joint distributions of random variables associated with fluctuations of a process with independent increments. *Theory of Probability and its Applications* **14**, 410–423.
- [27] PISTORIUS, M. (2006). On maxima and ladder processes for a dense class of Lévy process. *Journal of Applied Probability* **43**, 208–220.
- [28] SCHAASBERGER, R. (1973). *Warteschlangen*. Wien-New York: Springer-Verlag.