

# Workshop on Parameter Redundancy Part I

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Byron Morgan  
University of Kent

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# Collaborators

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- ▣ Ted Catchpole (Canberra and Kent)
- ▣ Diana Cole (Kent)

# Outline

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- Introduction and motivation
- Definitions
- General rules
- Use of symbolic algebra
- Extension theorems
- Near redundancy
- Weak identifiability

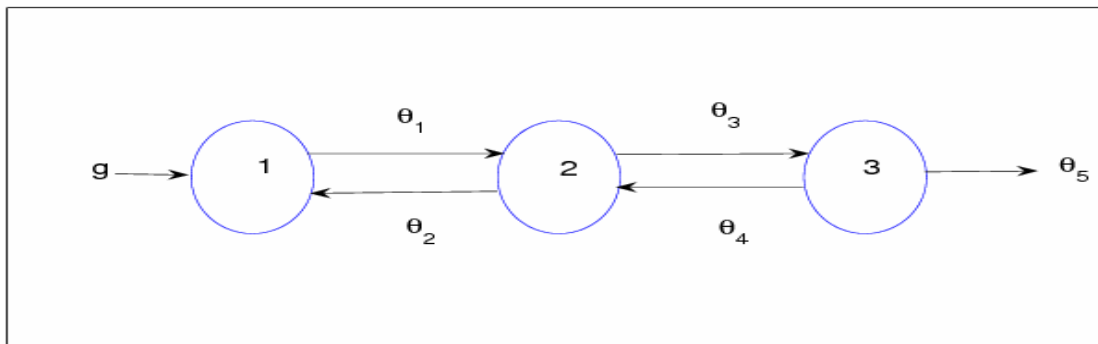
# Complex models and their parameters

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- Compartment models
- Ecology
- Econometrics
- Hidden Markov models

# Compartment models

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# Econometrics

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- Identifiability of the **simultaneous equation model**:

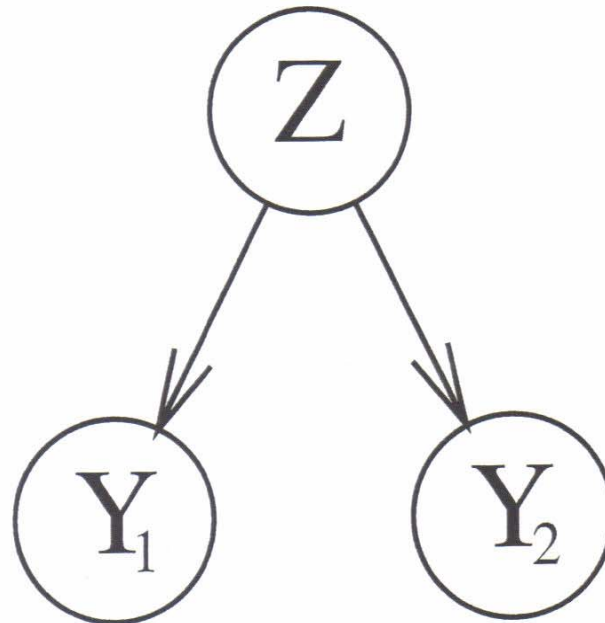
$$By_t + \Gamma x_t = u_t,$$

where  $y_t$  and  $u_t$  are vectors of random variables,  $x_t$  is a vector of non-random exogenous variables,  $B$  and  $\Gamma$  are matrices of parameters, and  $u_t$  has a normal distribution, with dispersion matrix  $\Sigma$ .

- The parameter space is  $[B, \Gamma, \Sigma]$ , some of which may be constrained.

# A simple naïve Bayesian network

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A naïve Bayesian network with a binary root node and two binary observable nodes.

# Ecology

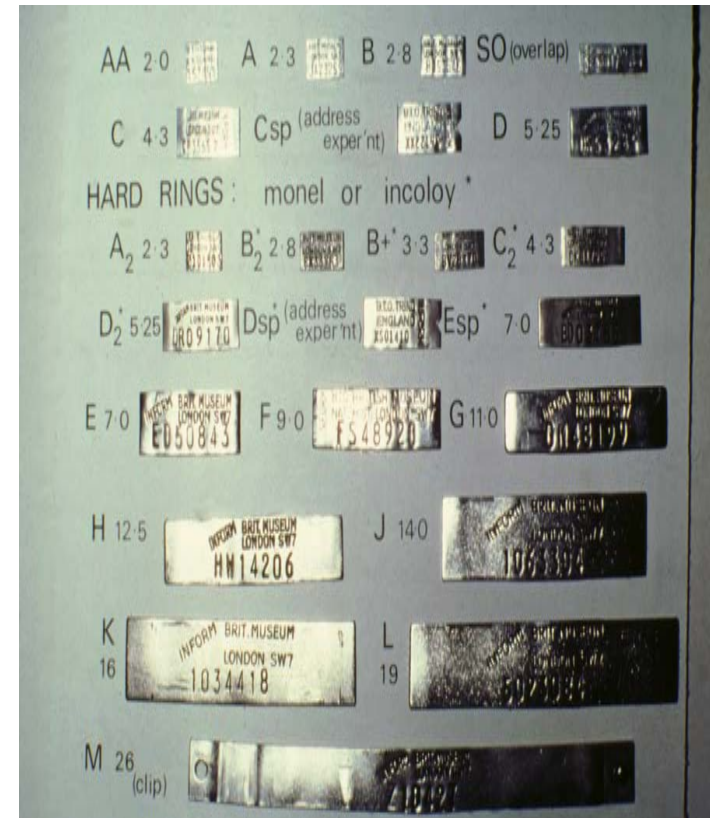
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- ❑ Estimation of the annual survival probabilities of wild animals.
- ❑ Collect data on previously marked animals.
- ❑ These are either found **dead or alive**.
- ❑ Form probability models.
- ❑ Fit to data using maximum likelihood, or Bayesian methods.



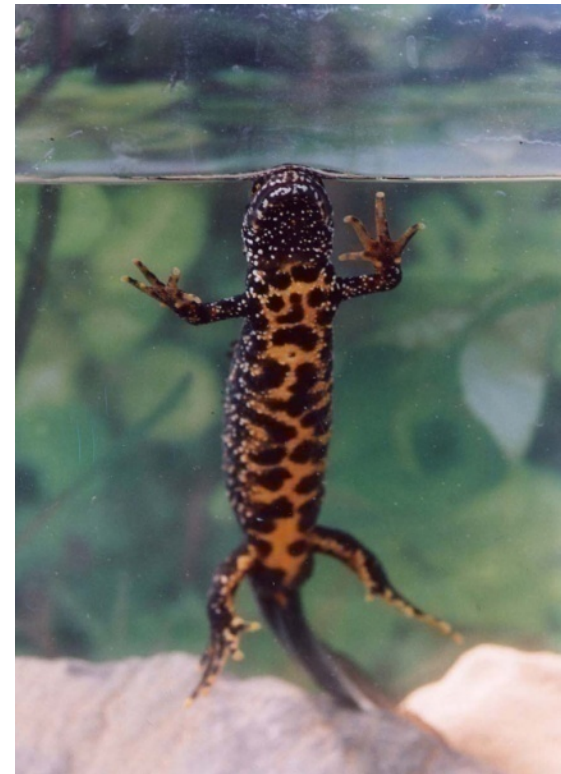
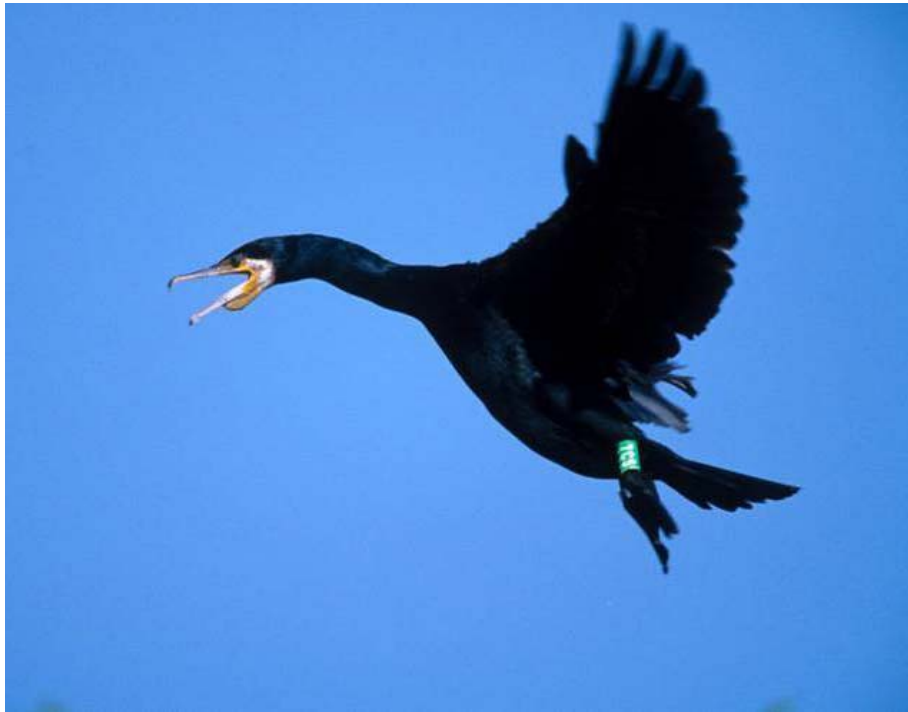
# Models for survival: Marking

- We obtain information on survival from studying previously marked animals
- These may be observed again alive or dead.
- It is assumed that marking does not affect behaviour



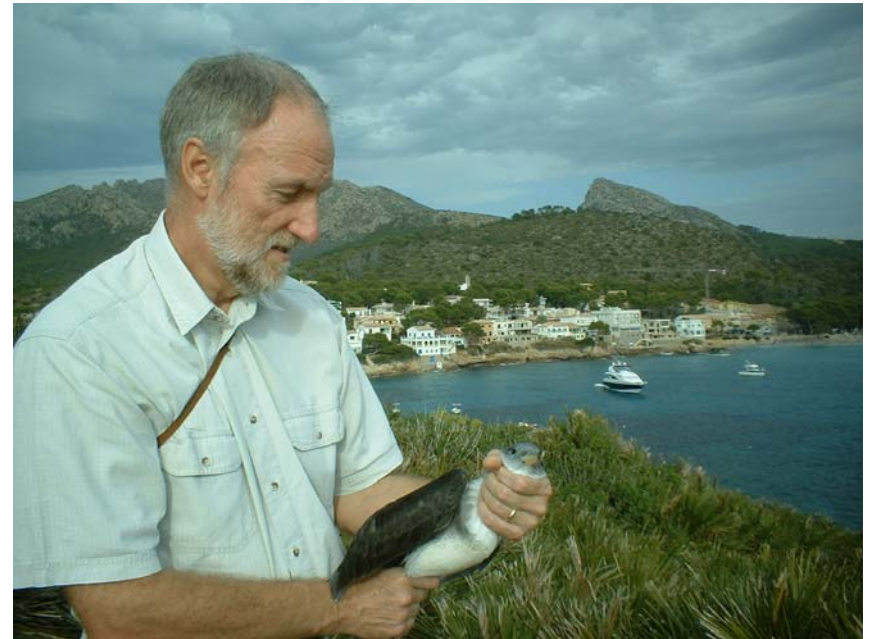
# Identification of Cormorant, *Phalacrocorax carbo sinensis*, and great crested newt, *Triturus cristatus*

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# Recapture of Cory's shearwater, *Calonectris diomedea*

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# Recovery/recapture

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Estimation of survival/mortality from information on the recovery of dead marked animals and from observations on live marked animals.





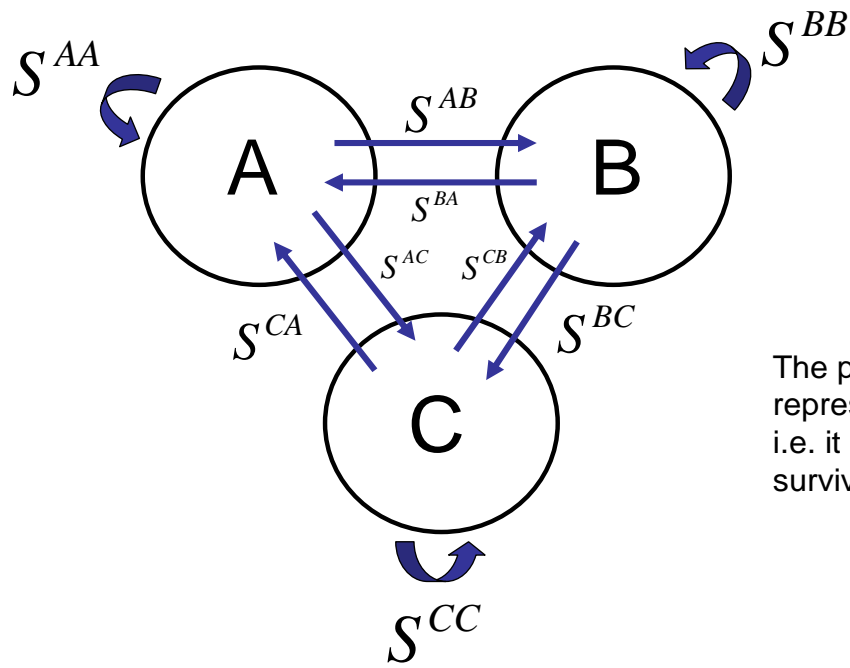
# Complexity

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- ❑ Models may be complicated, incorporating age, cohort and time components.
- ❑ Models may be simplified by the use of **covariates**.
- ❑ Modern focus on **multi-site data** can produce models with many parameters.
- ❑ **It is often unclear how many parameters can be estimated.**

# An example of a multi-site system

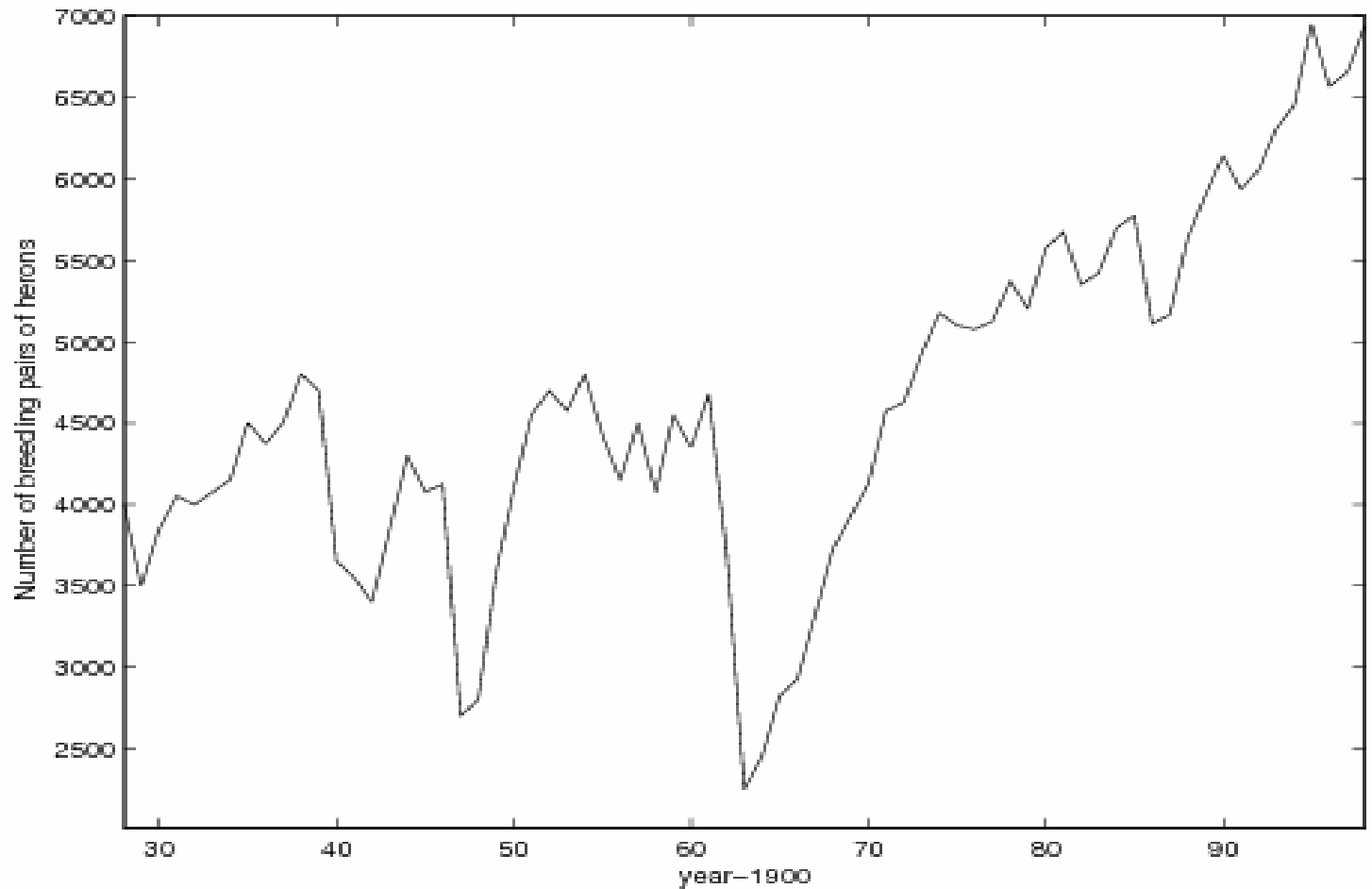
## Multisite Systems



The parameter  $S$  represents the “transition”, i.e. it represents both survival and movement

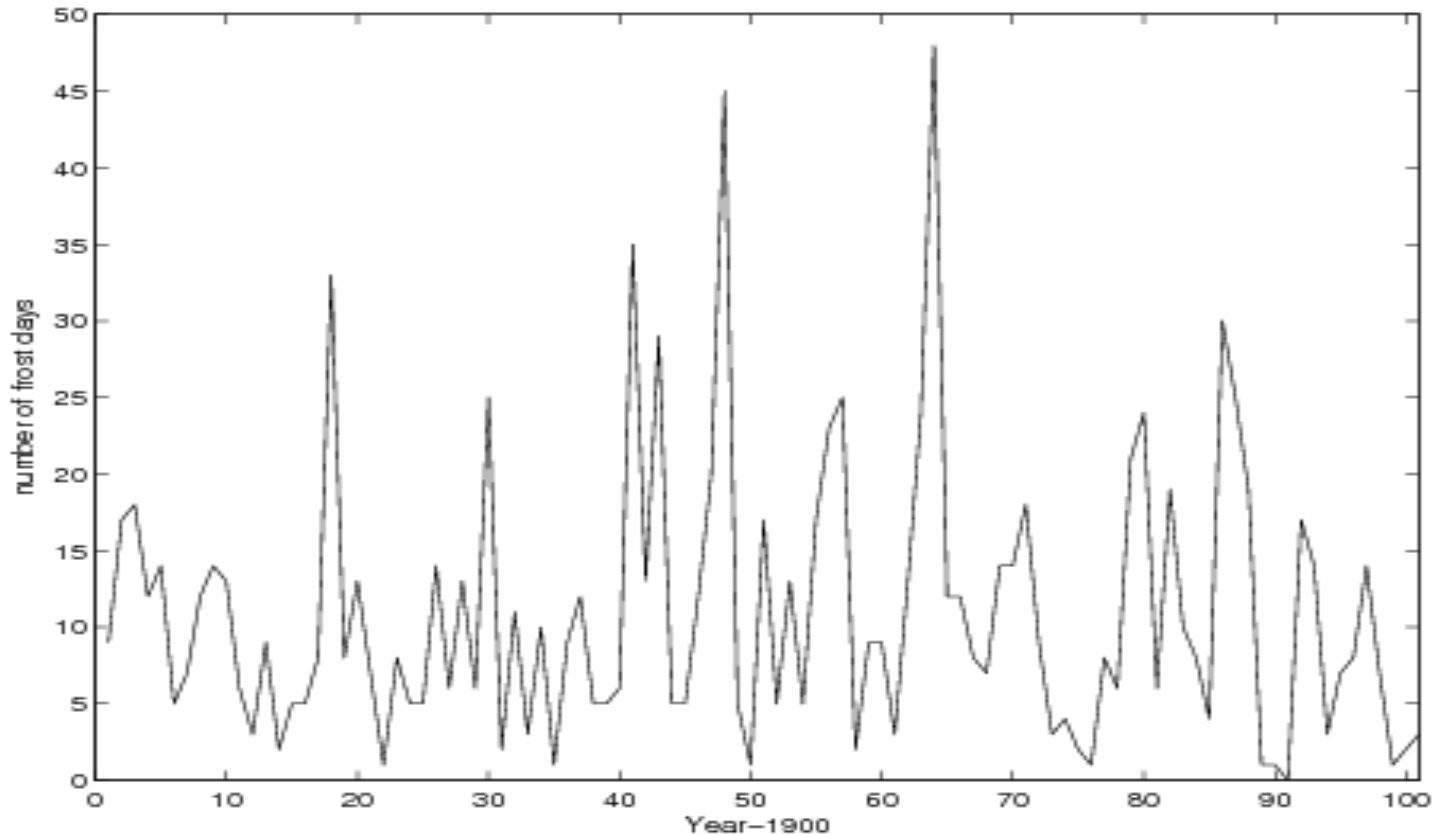
# The British heron census, *Ardea cinerea*

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# Climatic covariates: number of frost-days in **Central** England.

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# The Cormack-Jolly-Seber (CJS) model (1965)

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Consider a simple case in which all animals are adults, sharing a common probability of annual survival,  $\phi$ . If  $p$  denotes the probability of recapture then the **multinomial** probabilities corresponding to any cohort, of known size, of marked birds have the form:

$$\phi p, \quad \phi^2 p(1-p), \quad \phi^3 p(1-p)^2, \quad \dots$$

Parameters may be time-dependent – appropriate for adult animals.

# Illustration of CJS recapture probabilities: a 3-year study

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$\phi_1 p_2$	$\phi_1 \phi_2 (1-p_2)p_3$	$\phi_1 \phi_2 \phi_3 (1-p_2)(1-p_3)p_4$
	$\phi_2 p_3$	$\phi_2 \phi_3 (1-p_3)p_4$
		$\phi_3 p_4$

# CJS recapture probabilities: what we can estimate

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$\phi_1 p_2$	$\phi_1 \phi_2 (1-p_2)p_3$	$\phi_1 \phi_2 \phi_3 (1-p_2)(1-p_3)p_4$
	$\phi_2 p_3$	$\phi_2 \phi_3 (1-p_3)p_4$
		$\phi_3 p_4$

# Parameter redundancy

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- This model has deficiency of one: we can only estimate the product,  $\phi_3\rho_4$ . All the other parameters can be estimated.

# Parameter redundancy

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- This model has deficiency of one: we can only estimate the product,  $\phi_3 p_4$ . All the other parameters can be estimated.
  
- What if we only have two years of ringing?

# Illustration of CJS recapture probabilities: a 3-year study + 2 cohorts

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$\phi_1 p_2$	$\phi_1 \phi_2 (1-p_2)p_3$	$\phi_1 \phi_2 \phi_3 (1-p_2)(1-p_3)p_4$
	$\phi_2 p_3$	$\phi_2 \phi_3 (1-p_3)p_4$

# Illustration of CJS recapture probabilities: a 3-year study + 2 cohorts

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$\phi_1 p_2$	$\phi_1 \phi_2 (1-p_2)p_3$	$\phi_1 \phi_2 \phi_3 (1-p_2)(1-p_3)p_4$
	$\phi_2 p_3$	$\phi_2 \phi_3 (1-p_3)p_4$

# References

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- Goodman, 1974, *Biometrika*.
- Rothenberg, 1971, *Econometrica*.
- Walter, 1982, *Identifiability of state space models*.



# Section

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- Introduction and motivation
- **Definitions**
- General rules
- Use of symbolic algebra
- Expansion theorems
- Near redundancy
- Weak identifiability

# Parameter redundancy and identifiability

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- A model is **identifiable** if no two values of the parameters give the same probability distribution for the data.
- A model is **locally identifiable** if there is a distance  $\delta > 0$ , such that any two parameter values that give the same distribution must be separated by at least  $\delta$ .
- A **parameter redundant** model has parameters that cannot be estimated.
- A parameter redundant model is not locally identifiable.
- Full rank models are **essentially** or **conditionally** full rank.
- An essentially full rank model is locally identifiable.
- Are essentially full rank models identifiable?

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# General rules

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- ❑ In some cases it is possible to establish general rules for models of particular structures.
- ❑ This avoids having to use Maple (see later).
- ❑ A particular illustration of this occurs with **age-dependent** recovery models

# Model notation for recovery models

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Ring-recovery models are described as, for example:

C/A/C, T/A/C, T/A/T, C/C/T.

In this notation, each model is specified by 3 letters, which designate, in order,

1. The way we model **first-year survival**: C or T;
2. The way we model **adult survival**: C, A or T; and A can have categories.
3. The way we model the **recovery probability**: C, A or T.

We use this notation in **Example 4 of the first Practical**.

# Steps: age-dependence also in $\lambda$ .

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□ Consider, for example, the model denoted by C/A(2,2,3)/A(2,1,1,4). What can we estimate here?

□ Here we have the parameters:

$$\begin{array}{cccccccc} \phi_1, & \phi_2, & \phi_2, & \phi_3, & \phi_3, & \phi_4, & \phi_4, & \phi_4 \\ \lambda_1, & \lambda_1, & \lambda_2, & \lambda_3, & \lambda_4, & \lambda_4, & \lambda_4, & \lambda_4 \end{array}$$

# Steps: age-dependence also in $\lambda$ .

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- Consider, for example, the model denoted by C/A(2,2,3)/A(2,1,1,4). What can we estimate here?
- Here we have a **single step**, as shown:

$$\begin{array}{c} \phi_1, \phi_2, \phi_2 \mid \phi_3, \phi_3, \phi_4, \phi_4, \phi_4 \\ \lambda_1, \lambda_1, \lambda_2 \mid \lambda_3, \lambda_4, \lambda_4, \lambda_4, \lambda_4 \end{array}$$

# Theorem 1

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- Suppose the first step occurs at age  $n$ , and let  $m$  be the number of parameters used in the first  $n$  years.
- If  $m = n + 1$ , the model is parameter redundant.
- If  $1 < m < n + 1$ , then the step does not cause parameter redundancy. Furthermore, to test for parameter redundancy, the parameters occurring in the first  $n$  years can be discarded, and the count started anew in year  $n + 1$ .



# Theorem 2

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- In the age-dependent model T/A/A
- The step at age 1 year does **not** cause parameter-redundancy
- To determine any possible redundancy caused by a subsequent step, the age and parameter counts begin again after age 1 year, as in Theorem 1.

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# How to test for parameter redundancy in general

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- Form an appropriate **derivative matrix**,  $D$ .
- Use **Maple** to determine the **symbolic row rank** of  $D$ . Use this to determine if the model is parameter redundant or full rank.
- We can also determine which **parameter combinations** can be estimated, if the model is parameter redundant.

# The method

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The approach was for **exponential family** models. It is performed using a symbolic algebra package such as Maple.

1. Calculate  $\mathbf{D} = \left[ \frac{\partial \mu_j}{\partial \theta_i} \right]$  ( $\mu$  is the mean,  $\theta$  are parameters).

2. The number of estimable parameters =  $\text{rank}(\mathbf{D})$ .

3. Solve  $\alpha^T \mathbf{D} = 0$ . The location of the zeros in  $\alpha$  indicates which are the estimable parameters.

4. Solve  $\sum_{i=1}^p \alpha_{ij} \frac{\partial f}{\partial \theta_i} = 0$  to find the full set of estimable

parameters; ( $j$  is the index for  $>1$  solution to  $\alpha^T \mathbf{D} = 0$ ).

# Example 1: Cormack-Jolly-Seber Model

Little Penguins, *Eudyptula minor*, capture recapture data (1994 to 1997)

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$$\mathbf{N} = \begin{bmatrix} 30 & 58 & 37 \\ 0 & 20 & 37 \\ 0 & 0 & 18 \end{bmatrix}$$

$\phi_i$  – probability a penguin survives from occasion  $i$  to  $i+1$

$p_i$  – probability a penguin is recaptured on occasion  $i$

The set of parameters is:  $\theta = [\phi_1, \phi_2, \phi_3, p_2, p_3, p_4]$



$$\mathbf{P} = \begin{bmatrix} \phi_1 p_2 & \phi_1 \bar{p}_2 \phi_2 p_3 & \phi_1 \bar{p}_2 \phi_2 \bar{p}_3 \phi_3 p_4 \\ 0 & \phi_2 p_3 & \phi_2 \bar{p}_3 \phi_3 p_4 \\ 0 & 0 & \phi_3 p_4 \end{bmatrix}$$

$$\bar{p}_2 = 1 - p_2 \text{ etc}$$

We can now use  $\mathbf{P}$ .

## Forming the derivative matrix (take logs first)

$$\mathbf{D} = \frac{\partial \ln(\mathbf{P})}{\partial \theta} = \begin{bmatrix} \phi_1^{-1} & \phi_1^{-1} & \phi_1^{-1} & 0 & 0 & 0 \\ 0 & \phi_2^{-1} & \phi_2^{-1} & \phi_2^{-1} & \phi_2^{-1} & 0 \\ 0 & 0 & \phi_3^{-1} & 0 & \phi_3^{-1} & \phi_3^{-1} \\ p_2^{-1} & -\overline{p_2}^{-1} & -\overline{p_2}^{-1} & 0 & 0 & 0 \\ 0 & p_3^{-1} & -\overline{p_3}^{-1} & p_3^{-1} & -\overline{p_3}^{-1} & 0 \\ 0 & 0 & p_4^{-1} & 0 & p_4^{-1} & p_4^{-1} \end{bmatrix}$$

$\text{rank}(\mathbf{D}) = 5 < 6$ , so the model is parameter redundant.

In order to see which of the original parameters we can estimate:

Set  $\alpha^T \mathbf{D} = 0 \Rightarrow \alpha^T = [0, 0, -\phi_3 / p_4, 0, 0, 1]$

Solving PDE, we find that the estimable parameters are:  $\phi_1, \phi_2, p_2, p_3, \phi_3 p_4$

# Example – Cormack-Jolly-Seber Model

## with covariates

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We now set

$$\phi_i = 1/\{1 + \exp(a + bx_i)\}$$



For example,  $x_i$  could be the mean annual banding weight, or the SOI.

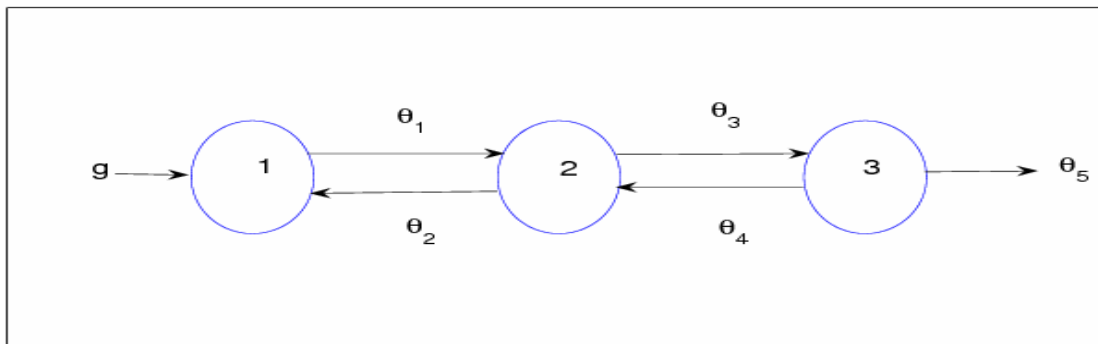
$$\theta = [a, b, \rho_2, \rho_3, \rho_4],$$

and we find that **the model is now full rank.**

See Maple practical.

# Compartment models

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# Simple compartment model

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$$\frac{dx_1}{dt} = -(\theta_1 + \theta_2)x_1 + \theta_3x_2 + u$$

$$\frac{dx_2}{dt} = \theta_2x_1 - (\theta_3 + \theta_4)x_2$$

$$y = x_1$$

Transfer function

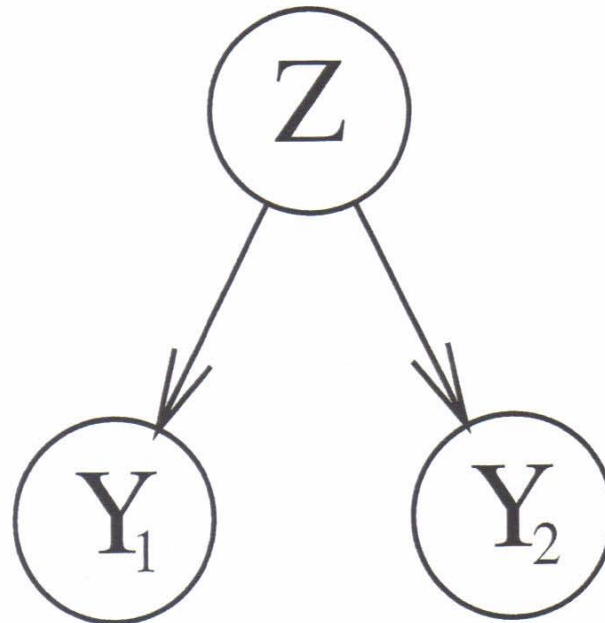
$$Q(s) = \frac{s + \theta_3 + \theta_4}{s^2 + s(\theta_1 + \theta_2 + \theta_3 + \theta_4) + \theta_1\theta_3 + \theta_1\theta_4 + \theta_2\theta_4}$$

Estimable parameters are:

$$(\theta_1 + \theta_2), -\theta_2\theta_3, (\theta_3 + \theta_4).$$

# A simple naïve Bayesian network

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A naïve Bayesian network with a binary root node and two binary observable nodes.

# Naïve Bayesian Networks in general

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- We have  $n$  observable nodes,  $Y_1, \dots, Y_n$ , and a single observable node  $Z$ .
- All nodes are binary.
- $2n+1$  parameters:  $p, \theta_{1|1}, \dots, \theta_{n|1}, \theta_{1|0}, \dots, \theta_{n|0}$ .

$$P(y) = p \prod_{i=1}^n \theta_{i|1}^{y_i} (1 - \theta_{i|1})^{1 - y_i} + (1 - p) \prod_{i=1}^n \theta_{i|0}^{y_i} (1 - \theta_{i|0})^{1 - y_i}$$

# Naïve Bayesian network ctd

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- In this example we can use a **reparameterisation** to show that
- For  $n > 2$  the model is full rank
- We can use the **PLUR decomposition** to determine parameter redundant sub-models: for example, when  $n = 3$ ,  
$$\text{Det}(U) = -p^3(1-p)^3(\theta_{1|1} - \theta_{1|0})^2(\theta_{2|1} - \theta_{2|0})^2(\theta_{3|1} - \theta_{3|0})^2.$$
- Previously conclusions followed a particular analysis.

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- **Extension theorems**
- Near redundancy
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# Extension theorems

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These give conditions which ensure that results which hold for a particular configuration also hold for **larger** configurations.

For instance, the CJS model **always** has deficiency one, when the number of cohorts equals the number of years of recapture.

# Extension Theorem

- Suppose a model has exhaustive summary  $\kappa_1$  and parameters  $\theta_1$ .

$$\mathbf{D}_1 = \left[ \frac{\partial \kappa_{1,j}}{\partial \theta_{1,i}} \right]$$

- Now extend that model by adding extra exhaustive summary terms  $\kappa_2$ , and extra parameters  $\theta_2$ . (Eg add more years of ringing/recovery) New model's exhaustive summary is  $\kappa = [\kappa_1 \ \kappa_2]^T$  and parameters are  $\theta = [\theta_1 \ \theta_2]^T$ .

$$\mathbf{D} = \begin{bmatrix} \left[ \frac{\partial \kappa_{1,j}}{\partial \theta_{1,i}} \right] & \left[ \frac{\partial \kappa_{2,j}}{\partial \theta_{1,i}} \right] \\ 0 & \left[ \frac{\partial \kappa_{2,j}}{\partial \theta_{2,i}} \right] \end{bmatrix} = \mathbf{D} = \begin{bmatrix} \mathbf{D}_1 & \left[ \frac{\partial \kappa_{2,j}}{\partial \theta_{1,i}} \right] \\ 0 & \mathbf{D}_2 \end{bmatrix}$$

- If  $\mathbf{D}_1$  is full rank and  $\mathbf{D}_2$  is full rank, the extended model will be full rank. The result can be further generalised by induction.
- Result is trivially always true, if you add zero or one extra parameters
- Method can also be used for parameter redundant models by first rewriting the model in terms of its estimable set of parameters.

# Extension

- Example: Ring-recovery model  $\theta = [\phi_{1,1} \ \phi_{1,2} \ \phi_a \ \lambda_1 \ \lambda_a]$  is a trivial example of extension theorem. Adding an extra year of ringing, adds one parameter. Adding an extra year of recovery adds no extra parameters.  $\therefore$  by induction general model rank is always full rank.

- Example: Ring-recovery Model (T/A/T) 4 years ringing 5 years recovery ( $p = 13$ )

$$\mathbf{P} = \begin{bmatrix} (1-\phi_{1,1})\lambda_1 & \phi_{1,1}(1-\phi_2)\lambda_2 & \phi_{1,1}\phi_2(1-\phi_3)\lambda_3 & \phi_{1,1}\phi_2\phi_3(1-\phi_4)\lambda_4 & \phi_{1,1}\phi_2\phi_3\phi_4(1-\phi_5)\lambda_5 \\ 0 & (1-\phi_{1,2})\lambda_2 & \phi_{1,2}(1-\phi_2)\lambda_3 & \phi_{1,2}\phi_2(1-\phi_3)\lambda_4 & \phi_{1,2}\phi_2\phi_3(1-\phi_4)\lambda_5 \\ 0 & 0 & (1-\phi_{1,3})\lambda_3 & \phi_{1,3}(1-\phi_2)\lambda_4 & \phi_{1,3}\phi_2(1-\phi_3)\lambda_5 \\ 0 & 0 & 0 & (1-\phi_{1,4})\lambda_4 & \phi_{1,4}(1-\phi_2)\lambda_5 \end{bmatrix}$$

Rank( $\mathbf{D}_1$ ) = 13,  $\therefore$  is full rank

- Add an extra year of recovery:

$$\mathbf{P} = \begin{bmatrix} (1-\phi_{1,1})\lambda_1 & \phi_{1,1}(1-\phi_2)\lambda_2 & \phi_{1,1}\phi_2(1-\phi_3)\lambda_3 & \phi_{1,1}\phi_2\phi_3(1-\phi_4)\lambda_4 & \phi_{1,1}\phi_2\phi_3\phi_4(1-\phi_5)\lambda_5 & \phi_{1,1}\phi_2\phi_3\phi_4\phi_5(1-\phi_6)\lambda_6 \\ 0 & (1-\phi_{1,2})\lambda_2 & \phi_{1,2}(1-\phi_2)\lambda_3 & \phi_{1,2}\phi_2(1-\phi_3)\lambda_4 & \phi_{1,2}\phi_2\phi_3(1-\phi_4)\lambda_5 & \phi_{1,2}\phi_2\phi_3\phi_4(1-\phi_5)\lambda_6 \\ 0 & 0 & (1-\phi_{1,3})\lambda_3 & \phi_{1,3}(1-\phi_2)\lambda_4 & \phi_{1,3}\phi_2(1-\phi_3)\lambda_5 & \phi_{1,3}\phi_2\phi_3(1-\phi_4)\lambda_6 \\ 0 & 0 & 0 & (1-\phi_{1,4})\lambda_4 & \phi_{1,4}(1-\phi_2)\lambda_5 & \phi_{1,4}\phi_2(1-\phi_3)\lambda_6 \end{bmatrix}$$

$\theta_2 = [\phi_6 \ \lambda_6]$ . Rank( $\mathbf{D}_2$ ) = 2,  $\therefore$  is full rank,  $\therefore$  extended model is full rank.

- Add an extra year of ringing, adds one parameter  $\therefore$  extended model is full rank. By induction model is T/A/T model is always full rank.



# Section

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- Data & near redundancy
- Weak identifiability

# Recapture of Dippers, *Cinclus cinclus*

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The table shows capture-recapture data for European Dippers in 1981-1986.

1981	22	11	2	0	0	0	0
1982	60		24	1	0	0	0
1983	78			34	2	0	0
1984	80				45	1	2
1985	88					51	0
1986	98						52

# Dealing with missing data & near-redundancy

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- ❑ **Missing data** in any application can **change the parameter redundancy status**.
- ❑ This is easily dealt with by removing the probabilities associated with empty cells.
- ❑ **Near-redundant** models are full-rank, but for certain data can result in poor estimation. This may be due to similarity to a parameter-redundant sub-model. Check **eigen-values of Hessian**.

# Section

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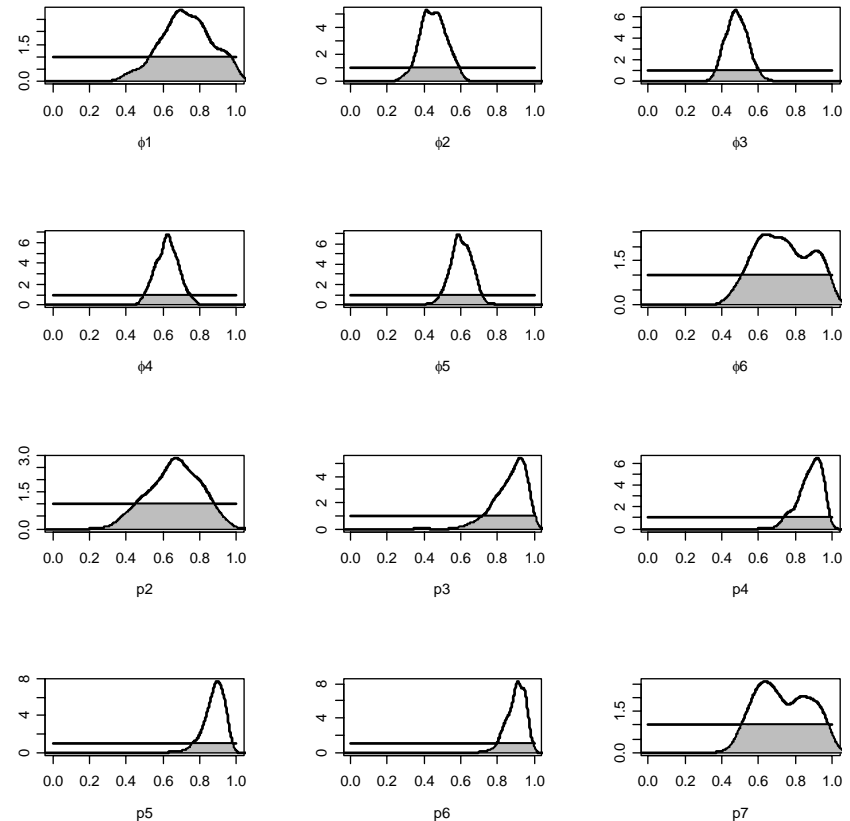
# Weak identifiability: the Bayesian context

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- A parameter  $\theta$  is said to be **weakly identifiable** when  $\pi(\theta|Y) \approx p(\theta)$ .
- This is the counterpart to near-redundancy.
- For each parameter in a model, Garrett and Zeger(2000) considered the overlap of prior and posterior.
- Form  $\tau = \int \min(p(\theta), \pi(\theta|Y))d\theta$ .
- Garrett and Zeger suggest ad-hoc threshold of  $\tau = 0.35$ . This works well for ecological applications.

# A Bayesian perspective: the CJS model

- In population ecology we may devise models with parameters that cannot be estimated from the data.
- Symbolic algebra can be used to examine whether a model is parameter-redundant.
- In a Bayesian context, it is interesting to consider the overlap between priors,  $p(\theta)$  and posteriors  $\pi(\theta|x)$ .



# Recapture of Dippers, *Cinclus cinclus*

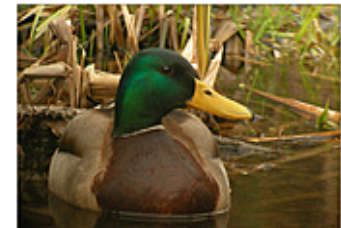
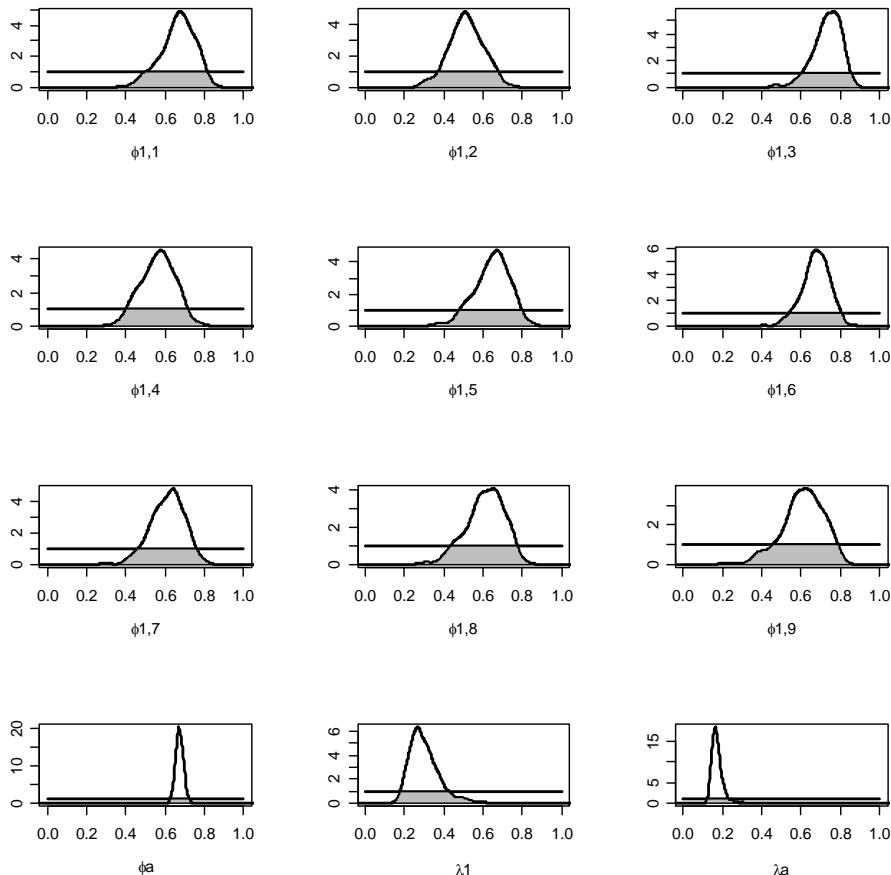
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Note small cohort size in 1981.

1981	22	11	2	0	0	0	0
1982	60		24	1	0	0	0
1983	78			34	2	0	0
1984	80				45	1	2
1985	88					51	0
1086	98						52

# Male mallard, *Anas platyrhynchos*

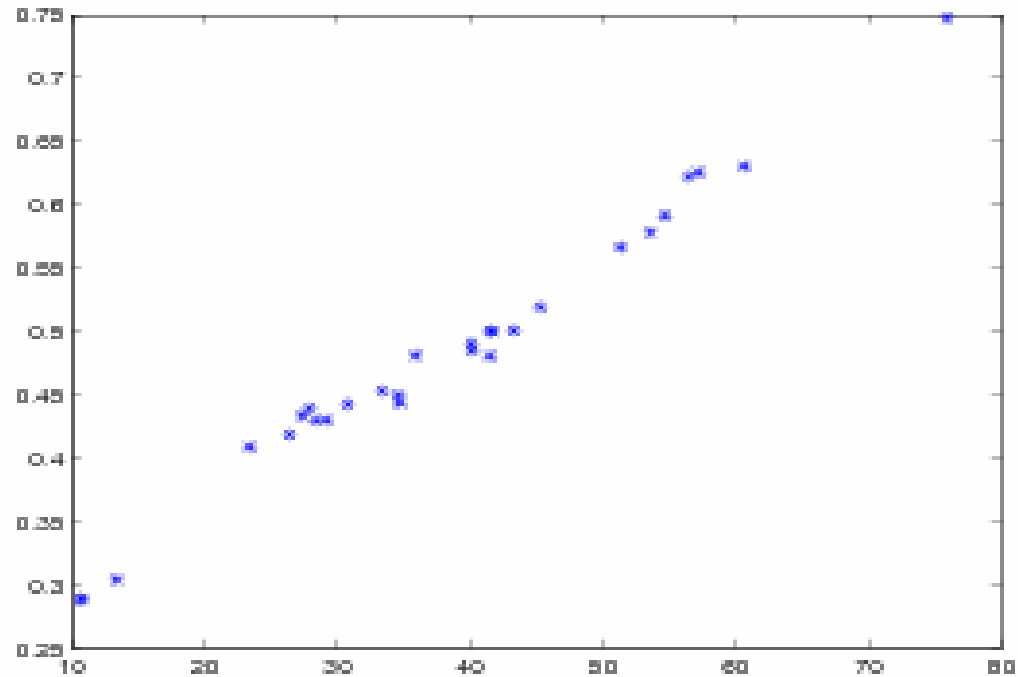


Model: T/C/A<sub>(1,1)</sub>  
 $\phi_{1,i}$ ,  $\phi_a$ ,  $\lambda_1$ ,  $\lambda_a$  here  
only two  
parameters,  $\phi_a$  and  
 $\lambda_a$  are strongly  
identified. The  
model is near-  
redundant.



# Relationship of overlap to interquartile range: simpler to calculate

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# Acknowledgement

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